

# THE TIME OF ULTIMATE RECOVERY IN GAUSSIAN RISK MODEL

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**Abstract:** We analyze the distance  $\mathcal{R}_T(u)$  between the first and the last passage time of  $\{X(t) - ct : t \in [0, T]\}$  at level  $u$  in time horizon  $T \in (0, \infty]$ , where  $X$  is a centered Gaussian process with stationary increments and  $c \in \mathbb{R}$ , given that the first passage time occurred before  $T$ . Under some tractable assumptions on  $X$ , we find  $\Delta(u)$  and  $G(x)$  such that

$$\lim_{u \rightarrow \infty} \mathbb{P}(\mathcal{R}_T(u) > \Delta(u)x) = G(x),$$

for  $x \geq 0$ . We distinguish two scenarios:  $T < \infty$  and  $T = \infty$ , that lead to qualitatively different asymptotics. The obtained results provide exact asymptotics of the ultimate recovery time after the ruin in Gaussian risk model.

**Key Words:** Gaussian risk process; exact asymptotics; first ruin time; last ruin time; generalized Pickands-Piterbarg constant.

**AMS Classification:** Primary 60G15; secondary 60G70, 60K25

## 1. INTRODUCTION

For given threshold  $u > 0$  and time horizon  $T \in (0, \infty]$ , let

$$\tau_T(u) := \inf \{t \geq 0 : X(t) - ct > u, 0 \leq t \leq T\}$$

and

$$\mathcal{T}_T(u) := \sup \{t \geq 0 : X(s) - cs > u, 0 \leq t \leq T\},$$

be *the first* and *the last passage time* of process  $X(t) - ct, t \geq 0$  at level  $u$  respectively, with convention that  $\inf \emptyset = \infty$  and  $\sup \emptyset = 0$ .

The analysis of properties of  $\tau_T(u)$  and  $\mathcal{T}_T(u)$ , due to their obvious importance in extreme value theory of stochastic processes, attracted substantial interest, being additionally stimulated by relations of passage times with important problems in applied probability. More specifically, in risk theory  $\tau_T(u)$  and  $\mathcal{T}_T(u)$  have the interpretation as *the first* and *the last ruin time* of the risk reserve process  $S(t) := u + ct - X(t)$ , where  $u > 0$  is the initial capital,  $c > 0$  is the premium rate and  $X(t), t \geq 0$  is the accumulated claim amount in interval  $(0, t]$ ; see e.g. [3, 18].

In this contribution we suppose that  $X$  is a centered Gaussian process with stationary increments, a.s. continuous sample paths,  $X(0) = 0$  and  $c > 0$ . We note that in the context of risk theory, there are strong application-based and theoretical reasons for modelling accumulated claim amount by Gaussian processes with stationary increments. On one hand the family of Gaussian processes provides flexibility in the adjustment of suitable correlation model, since it covers wide range of correlation structures. On the other hand, there are theoretical results that legitimate approximation of

the accumulated claim amount in highly aggregated models by Gaussian processes; see the celebrated work by Iglehart [22] for the Brownian approximation and, e.g., [7, 18, 26] for more general models, including e.g. fractional Brownian motion approximations.

This contribution is devoted to analysis of the distribution of

$$\mathcal{R}_T(u) := \mathcal{T}_T^*(u) - \tau_T^*(u),$$

where

$$(1) \quad (\tau_T^*(u), \mathcal{T}_T^*(u)) := (\tau_T(u), \mathcal{T}_T(u)) \Big|_{(\tau_T(u) < \infty)},$$

both for  $T \in (0, \infty)$  and  $T = \infty$ . Referring again to risk theory,  $\mathcal{R}_T(u)$  has the interpretation as *the ultimate time to recovery*, which is the difference between the last and first ruin time, under the condition that ruin occurred; see also [14, 23] and references therein. We note that  $\mathcal{R}_T(u)$  is also closely related to the so-called *Parisian ruin time*, which is the first time that the length of the consecutive excursion period of the surplus process  $S$  under level 0 exceeds a pre-specified time threshold; see, e.g., [2], [25], [4] and [5], with straightforward observation that  $\mathcal{R}_T(u)$  gives an upper bound for the appropriately chosen pre-specified time period in Parisian model. Another related notion is *the cumulative ruin time*, which is based on the total time spent below 0 (*in red*) by the underlying risk process; see, e.g., [15]. Clearly,  $\mathcal{R}_T(u)$  is greater than the corresponding occupation time.

For  $T = \infty$ , the asymptotics of the distribution of conditional first and last ruin times in Gaussian risk context were studied in, e.g., [20] and [21]; see also [18], [24] and [7] for related  $\gamma$ -reflected Gaussian models. Specifically, under some tractable assumptions on  $X$ , the following asymptotics was found in [7, 20, 21]:

$$(2) \quad \left( \frac{\tau_\infty^*(u) - ut_u}{\sigma(ut_u)}, \frac{\mathcal{T}_\infty^*(u) - ut_u}{\sigma(ut_u)} \right) \xrightarrow{d} (\mathcal{CN}, \mathcal{CN}), \quad u \rightarrow \infty,$$

where  $\sigma^2(t) = \text{Var}(X(t))$ ,  $t_u = \arg \inf_{t>0} \frac{u(1+ct)}{\sigma(ut)}$ ,  $\mathcal{N} \sim N(0, 1)$  and  $C$  is some known constant. However, the above result is too crude in order to deal with  $\mathcal{R}_\infty(u)$ , since it follows straightforwardly from (2), that

$$\frac{\mathcal{T}_\infty^*(u) - \tau_\infty^*(u)}{\sigma(ut_u)} \rightarrow 0, \quad u \rightarrow \infty$$

in probability; see also Corollary 4 in [21]. This implies that  $\sigma(ut_u)$  acting as a scaling function is too big, so in order to get a nontrivial result one has to scale by asymptotically smaller function.

The main results of this contribution provide  $\Delta(u)$  and distribution function  $G(\cdot)$  such that

$$\lim_{u \rightarrow \infty} \mathbb{P}(\mathcal{R}_T(u) > x\Delta(u)) = G(x),$$

for  $x \geq 0$ , both for  $T \in (0, \infty)$  and  $T = \infty$ .

As it is shown in Theorems 2.1, 2.3, both  $\Delta(u)$  and  $G(x)$  depend on  $T$  and on the local behavior of variance function of  $X$ , which leads to several scenarios. Interestingly, the limit function  $G$  is given in terms of *generalized Pickands-Piterebarg-type* constants. In order to obtain the main results of this contribution we accommodate to our needs recently developed *uniform double-sum method* applied for relevant continuous functionals; see [6].

Organization of the paper: Section 2 is devoted to introduction of notation and presentation of main results. In section 3 we present proofs of the main results.

## 2. MAIN RESULTS

In this section we provide main results of this contribution, which is the limit theorem for

$$\mathcal{R}_T(u) = \mathcal{T}_T^*(u) - \tau_T^*(u),$$

as  $u \rightarrow \infty$ , for  $T \in (0, \infty]$ . Due to specific asymptotic nature of  $\mathcal{R}_T(u)$  we distinguish two separate scenarios: infinite-time horizon ( $T = \infty$ ) and finite-time horizon ( $T < \infty$ ).

**2.1. Infinite-time horizon.** Suppose that  $T = \infty$ . Consider a centered Gaussian process  $X$  with continuous trajectories, stationary increments and variance function  $\sigma^2(t) := \text{Var}(X(t))$  such that **AI**:  $\sigma^2(0) = 0$ ,  $\sigma^2(t)$  is regularly varying at  $\infty$  with index  $2\alpha_\infty \in (0, 2)$  and  $\sigma^2(t)$  is twice continuously differentiable on  $(0, \infty)$  with its first derivative  $\dot{\sigma}^2(t) := \frac{d\sigma^2}{dt}(t)$  and second derivative  $\ddot{\sigma}^2(t) := \frac{d^2\sigma^2}{dt^2}(t)$  being ultimately monotone at  $\infty$ .

**AII**:  $\sigma^2(t)$  is regularly varying at 0 with index  $2\alpha_0 \in (0, 2]$ .

For given  $x \geq 0$  and  $f \in C([0, 2S])$ ,  $S > 0$ , let

$$(3) \quad \Gamma(x, S; f) = \sup_{t \in [0, S]} \min \left( f(t), \sup_{s \in [t+x, 2S]} f(s) \right).$$

Then, for given  $\eta(t), t \geq 0$ , a centered Gaussian process with stationary increments and continuous sample paths, we introduce

$$(4) \quad \begin{aligned} \mathcal{H}_\eta^\Gamma(x, S) &:= \mathbb{E} \left( \exp \left( \Gamma(x, S; \sqrt{2}\eta(t) - \text{Var}(\eta(t))) \right) \right), \\ \mathcal{H}_\eta^\Gamma(x) &:= \lim_{S \rightarrow \infty} \frac{\mathcal{H}_\eta^\Gamma(x, S)}{S}, \end{aligned}$$

providing that the limit exists. We note that  $\mathcal{H}_\eta^\Gamma(0)$  coincides with the notion of *generalized Pickands constant*, since

$$\mathcal{H}_\eta^\Gamma(0, S) = \mathbb{E} \left( \exp \left( \sup_{t \in [0, S]} (\sqrt{2}\eta(t) - \text{Var}(\eta(t))) \right) \right);$$

see [27], [10], [19], [11] [8], [16], [13], [12] and [17] for properties of generalized Pickands constants. In order to simplify notation, let

$$(5) \quad \mathcal{H}_\eta := \mathcal{H}_\eta^\Gamma(0), \quad \text{and} \quad \mathcal{H}_\eta([0, S]) := \mathcal{H}_\eta^\Gamma(0, S).$$

Let  $B_H(t)$  denote the standard fractional Brownian motion with mean 0 and correlation function satisfying

$$\text{Cov}(B_H(s), B_H(t)) = \frac{|s|^{2H} + |t|^{2H} - |t-s|^{2H}}{2}, \quad s, t \geq 0, H \in (0, 1].$$

Let  $t^* = \frac{\alpha_\infty}{c(1-\alpha_\infty)}$  and  $\overleftarrow{\sigma}(t), t \geq 0$  stand for the asymptotic inverse function of  $\sigma$  at value of  $t$ . Furthermore, let

$$(6) \quad \Delta(u) = \overleftarrow{\sigma} \left( \frac{\sqrt{2}\sigma^2(ut^*)}{u(1+ct^*)} \right).$$

In the rest of this section we tacitly assume that

$$\varphi := \lim_{u \rightarrow \infty} \frac{\sigma^2(u)}{u} \in [0, \infty].$$

**Theorem 2.1.** *Let  $X(t)$  be a centered Gaussian process with continuous trajectories and stationary increments satisfying **AI-AII**. Then for any  $x \geq 0$*

$$\lim_{u \rightarrow \infty} \mathbb{P} \left( \frac{\mathcal{R}_\infty(u)}{\Delta(u)} > x \right) = \frac{\mathcal{H}_\eta^\Gamma(x)}{\mathcal{H}_\eta^\Gamma(0)} \in (0, 1],$$

where

$$(7) \quad \eta(t) = \begin{cases} B_{\alpha_0}(t), & \varphi = 0 \\ \frac{X(\varphi t)}{\sigma(\varphi)}, & \varphi \in (0, \infty) \\ B_{\alpha_\infty}(t), & \varphi = \infty. \end{cases}$$

**Remark 2.2.** *Due to the covariance decomposition in Lemma 3.2, process  $\eta(t)$  in the generalized Pickands-Piterbarg constant is determined by  $\lim_{u \rightarrow \infty} \frac{\sigma^2(\Delta(u)t)}{\sigma^2(\Delta(u))}$ . In particular, if  $\varphi = 0$ , then  $\lim_{u \rightarrow \infty} \Delta(u) = 0$  and thus, by the regularity of  $\sigma^2(t)$  at 0, it follows that  $\lim_{u \rightarrow \infty} \frac{\sigma^2(\Delta(u)t)}{\sigma^2(\Delta(u))} = t^{2\alpha_0}$ , which is the variance of fractional Brownian motion  $B_{2\alpha_0}(t)$ . The case  $\varphi = \infty$  leads to  $B_{2\alpha_\infty}(t)$  by the same argument. If  $\varphi \in (0, \infty)$ , then  $\lim_{u \rightarrow \infty} \Delta(u) \in (0, \infty)$ , implying that  $\lim_{u \rightarrow \infty} \frac{\sigma^2(\Delta(u)t)}{\sigma^2(\Delta(u))} = \frac{\sigma^2(\varphi t)}{\sigma^2(\varphi)}$ , which corresponds to the variance of  $\frac{X(\varphi t)}{\sigma(\varphi)}$ .*

**2.2. Finite-time horizon.** In this subsection we focus on the finite-time case, i.e. we suppose that  $T \in (0, \infty)$ . Let  $X(t), t \in [0, T]$  be a Gaussian process with stationary increments, a.s. continuous trajectories, zero-mean and variance function  $\sigma^2$  satisfying

**BI**  $\sigma^2(0) = 0$  and  $\sigma^2 \in C(0, T]$  with the first derivative being positive.

**BII**  $\sigma^2$  is regularly varying at 0 with index  $2\alpha_0 \in (0, 2]$ .

Denote by

$$\Delta_1(u) = \overleftarrow{\sigma} \left( \frac{\sqrt{2}\sigma^2(T)}{u + cT} \right), \quad \Delta_2(u) = \left( \frac{\sigma(T)}{u + cT} \right)^2,$$

and

$$(8) \quad \begin{aligned} \mathcal{P}_{B_{1/2}}^d(x) &= \int_0^\infty e^w \mathbb{P} \left( \sup_{t \in [x, \infty)} \min \left( \sup_{s \in [0, t-x]} \sqrt{2}B_{1/2}(s) - (1+d)s, \sqrt{2}B_{1/2}(t) - (1+d)t \right) > w \right) dw \\ &+ \int_{-\infty}^0 e^w \mathbb{P} \left( \sup_{t \in [x, \infty)} \left( \sqrt{2}B_{1/2}(t) - (1+d)t \right) > w \right) dw, \end{aligned}$$

with  $B_{1/2}$  a standard Brownian motion and  $d > 0$ . We note that

$$\mathcal{P}_{B_{1/2}}^d(0) = \mathbb{E} \left( \exp \left( \sup_{t \in [0, \infty)} \left( \sqrt{2}B_{1/2}(t) - (1+d)t \right) \right) \right),$$

is the classical Piterbarg constant (see [1] and reference therein) and hence

$$\mathcal{P}_{B_{1/2}}^d(0) = 1 + \frac{1}{d};$$

see, e.g., [27].

**Theorem 2.3.** *Let  $X(t), t \in [0, T]$  be a centered Gaussian process with continuous trajectories and stationary increments satisfying **BI-BII** and  $x \geq 0$ .*

i) *If  $t = o(\sigma^2(t)), t \rightarrow 0$ , then*

$$\lim_{u \rightarrow \infty} \mathbb{P} \left( \frac{\mathcal{R}_T(u)}{\Delta_1(u)} > x \right) = \frac{\mathcal{H}_{B_{\alpha_0}}^\Gamma(x)}{\mathcal{H}_{B_{\alpha_0}}^\Gamma(0)} \in (0, 1].$$

ii) *If  $\sigma^2(t) \sim at, t \rightarrow 0, a > 0$ , then for  $d = \frac{2\sigma(T)\dot{\sigma}(T)}{a}$ ,*

$$\lim_{u \rightarrow \infty} \mathbb{P} \left( \frac{\mathcal{R}_T(u)}{\Delta_1(u)} > x \right) = \frac{\mathcal{P}_{B_{1/2}}^d(x)}{\mathcal{P}_{B_{1/2}}^d(0)} \in (0, 1].$$

iii) *If  $\sigma^2(t) = o(t), t \rightarrow 0$ , then*

$$\lim_{u \rightarrow \infty} \mathbb{P} \left( \frac{\mathcal{R}_T(u)}{\Delta_2(u)} > x \right) = e^{-\frac{\dot{\sigma}(T)}{\sigma(T)}x}.$$

### 3. PROOFS

In this section, we give proofs of Theorem 2.1 and Theorem 2.3. Hereafter, denote by  $\mathbb{Q}, \mathbb{Q}_i, i = 1, 2, 3, \dots$  positive constants that may differ from line to line and  $\bar{X} := \frac{X}{\sqrt{\text{Var}(X)}}$  for any nontrivial random variable  $X$ . Moreover,  $f(u) \sim g(u), u \rightarrow \infty$  means that  $\lim_{u \rightarrow \infty} \frac{f(u)}{g(u)} = 1$ . In our proofs, multiple limits appear. We shall write for instance

$$b(u, S, \epsilon) \sim a(u), \quad u \rightarrow \infty, S \rightarrow \infty, \epsilon \rightarrow 0$$

to mean that

$$\lim_{\epsilon \rightarrow 0} \lim_{S \rightarrow \infty} \lim_{u \rightarrow \infty} \frac{b(u, S, \epsilon)}{a(u)} = 1.$$

**3.1. Infinite-time horizon.** Observe that for any  $x \geq 0$

$$(9) \quad \mathbb{P} \left( \frac{\mathcal{T}_\infty^*(u) - \tau_\infty^*(u)}{\Delta(u)} > x \right) = \frac{\mathbb{P}(\mathcal{T}_\infty(u) - \tau_\infty(u) > x\Delta(u), \tau_\infty < \infty)}{\mathbb{P}(\tau_\infty(u) < \infty)},$$

with  $\Delta(u)$  defined in (6). In order to derive the limiting distribution of the above ratio, we need to derive the asymptotics of  $\mathbb{P}(\mathcal{T}_\infty(u) - \tau_\infty(u) > x\Delta(u), \tau_\infty < \infty)$  and  $\mathbb{P}(\tau_\infty(u) < \infty)$  respectively. Using that

$$\mathbb{P}(\tau_\infty(u) < \infty) = \mathbb{P} \left( \sup_{t \geq 0} X(t) - ct > u \right) = \mathbb{P} \left( \sup_{t \geq 0} X_u(t) > m(u) \right),$$

where  $X_u(t) = \frac{X(ut)}{u(1+ct)}m(u)$  with  $m(u) = \inf_{t > 0} \frac{u(1+ct)}{\sigma(ut)}$ , Proposition 2 in [11] (or Theorems 3.1-3.3 in [9]), implies the following Lemma.

**Lemma 3.1.** *Let  $X(t)$  be a centered Gaussian process with continuous trajectories and stationary increments satisfying **AI-AII**. Assuming that  $\lim_{u \rightarrow \infty} \frac{\sigma^2(u)}{u} \in [0, \infty]$ , we have*

$$\mathbb{P}(\tau_\infty(u) < \infty) \sim \mathcal{H}_\eta^\Gamma(0) \sqrt{\frac{2A\pi}{B}} \frac{u}{m(u)\Delta(u)} \Psi(m(u)),$$

where  $\Delta(u)$  is defined in (6),  $\eta$  is defined in (7) and

$$A = \left( \frac{\alpha_\infty}{c(1-\alpha_\infty)} \right)^{-\alpha_\infty} \frac{1}{1-\alpha_\infty}, \quad B = \left( \frac{\alpha_\infty}{c(1-\alpha_\infty)} \right)^{-\alpha_\infty-2} \alpha_\infty.$$

Thus, by (9), we are left with finding the asymptotics of  $\mathbb{P}(\mathcal{T}_\infty(u) - \tau_\infty(u) > x\Delta(u), \tau_\infty < \infty)$ , as  $u \rightarrow \infty$ .

In the next lemma we focus on asymptotic properties of the variance and correlation functions of relevant Gaussian processes; we refer to, e.g., [9] for the proof. Let

$$\sigma_u(t) := \text{Var}^{1/2}(X_u(t)) = \text{Var}^{1/2}\left(\frac{X(ut)}{u(1+ct)}m(u)\right), t \geq 0, \quad t_u = \arg \inf_{t>0} \frac{u(1+ct)}{\sigma(ut)}.$$

**Lemma 3.2.** *Suppose that AI-AII are satisfied. For  $u$  large enough  $t_u$  is unique, and  $t_u \rightarrow t^* = \frac{\alpha_\infty}{c(1-\alpha_\infty)}$ , as  $u \rightarrow \infty$ . Moreover, for any  $\delta_u > 0$  with  $\lim_{u \rightarrow \infty} \delta_u = 0$*

$$\lim_{u \rightarrow \infty} \sup_{t \in (t_u - \delta_u, t_u + \delta_u) \setminus \{t_u\}} \left| \frac{1 - \sigma_u(t)}{\frac{B}{2A}(t - t_u)^2} - 1 \right| = 0,$$

and

$$\lim_{u \rightarrow \infty} \sup_{s \neq t, s, t \in (t_u - \delta_u, t_u + \delta_u)} \left| \frac{1 - \text{Cor}(X(us), X(ut))}{\frac{\sigma^2(u|s-t|)}{2\sigma^2(ut^*)}} - 1 \right| = 0.$$

**Proof of Theorem 2.1** Due to (9) and Lemma 3.1, we focus on the asymptotics of

$$\begin{aligned} & \mathbb{P}(\mathcal{T}_\infty(u) - \tau_\infty(u) > x\Delta(u), \tau_\infty < \infty) \\ & = \mathbb{P}(\exists s, t \geq 0, s - t \geq \Delta(u)x, X(t) - ct > u, X(s) - cs > u), \end{aligned}$$

as  $u \rightarrow \infty$ , for any  $x \geq 0$ . We have

$$(10) \quad \pi_1(u) \leq \mathbb{P}(\mathcal{T}_\infty(u) - \tau_\infty(u) > x\Delta(u), \tau_\infty < \infty) \leq \pi_1(u) + \pi_2(u),$$

where

$$\begin{aligned} \pi_1(u) &= \mathbb{P}(\exists s, t \in E_1(u), s - t \geq \Delta(u)x, X(t) - ct > u, X(s) - cs > u) \\ \pi_2(u) &= \mathbb{P}(\exists (s, t) \in [0, \infty)^2 \setminus E_1^2(u), s - t \geq \Delta(u)x, X(t) - ct > u, X(s) - cs > u), \end{aligned}$$

with

$$E_1(u) = \left[ ut_u - \frac{u \ln m(u)}{m(u)}, ut_u + \frac{u \ln m(u)}{m(u)} \right].$$

It follows that for  $u > 0$

$$\pi_2(u) \leq 2\mathbb{P}\left(\sup_{t \in [0, \infty) \setminus E_1(u)} X(t) - ct > u\right).$$

Hence, following Lemma 7 in [11] (or Lemma 5.6 in [9]), we have that

$$(11) \quad \pi_2(u) \leq 2\mathbb{P}\left(\sup_{t \in [0, \infty) \setminus E_1(u)} X(t) - ct > u\right) = o(\mathbb{P}(\tau_\infty(u) < \infty)), \quad u \rightarrow \infty.$$

Thus we are left with finding the exact asymptotics of  $\pi_1(u)$  as  $u \rightarrow \infty$ . Replacing  $t$  by  $ut_u + \Delta(u)t$  and  $s$  by  $ut_u + \Delta(u)s$ , we rewrite

$$(12) \quad \pi_1(u) = \mathbb{P}(\exists s, t \in E_2(u), s - t \geq x, Z_u(t) > m(u), Z_u(s) > m(u))$$

with

$$Z_u(t) = \frac{X(ut_u + \Delta(u)t)}{u(1+ct_u) + c\Delta(u)t}m(u), \quad E_2(u) = \left[ -\frac{u \ln m(u)}{\Delta(u)m(u)}, \frac{u \ln m(u)}{\Delta(u)m(u)} \right].$$

Bonferroni inequality gives that for  $S > x$ ,

$$(13) \quad \Sigma_1^+(u) - \Sigma\Sigma_1(u) \leq \pi_1(u) \leq \Sigma_1^-(u) + \Sigma\Sigma_2(u),$$

where

$$\begin{aligned} \Sigma_1^\pm(u) &= \sum_{k=-N(u)\pm 1}^{N(u)\mp 1} \mathbb{P}(\exists t \in [kS, (k+1)S], s \in [t+x, (k+2)S] : \min(Z_u(t), Z_u(s)) > m(u)) \\ \Sigma\Sigma_1(u) &= \sum_{k=-N(u)-1}^{N(u)+1} \sum_{l=k+1}^{N(u)+1} \mathbb{P}\left(\sup_{t \in [kS, (k+1)S]} Z_u(t) > m(u), \sup_{t \in [lS, (l+1)S]} Z_u(t) > m(u)\right) \\ \Sigma\Sigma_2(u) &= \sum_{k=-N(u)-1}^{N(u)+1} \sum_{l=k+2}^{N(u)+1} \mathbb{P}\left(\sup_{t \in [kS, (k+1)S]} Z_u(t) > m(u), \sup_{t \in [lS, (l+1)S]} Z_u(t) > m(u)\right), \end{aligned}$$

where  $N(u) = \left\lceil \frac{u \ln m(u)}{\Delta(u)m(u)S} \right\rceil$ . To get the asymptotics of  $\pi_1(u)$ , in next steps of the proof we show that  $\Sigma_1^+(u) \sim \Sigma_1^-(u)$  and  $\Sigma\Sigma_2(u) \leq \Sigma\Sigma_1(u) = o(\Sigma_1^+(u))$ , as  $u \rightarrow \infty$  and  $S \rightarrow \infty$ .

Asymptotics of  $\Sigma_1^\pm(u)$ . Setting

$$Z_{u,k}(t) = Z_u(kS + t),$$

we have that

$$\begin{aligned} \Sigma_1^-(u) &= \sum_{k=-N(u)-1}^{N(u)+1} \mathbb{P}\left(\sup_{t \in [kS, (k+1)S]} \min\left(Z_u(t), \sup_{s \in [t+x, (k+2)S]} Z_u(s)\right) > m(u)\right) \\ &= \sum_{k=-N(u)-1}^{N(u)+1} \mathbb{P}(\Gamma(x, S; Z_{u,k}) > m(u)) \\ &\leq \sum_{k=-N(u)-1}^{N(u)+1} \mathbb{P}\left(\Gamma(x, S; \bar{Z}_{u,k}) > \frac{m(u)}{\sup_{t \in [0, 2S]} \sqrt{\text{Var}(Z_{u,k}(t))}}\right), \end{aligned}$$

where  $\Gamma$  is defined in (3). By Lemma 3.2, for any  $0 < \epsilon < 1$

$$\frac{m(u)}{\sup_{t \in [kS, (k+1)S]} \sqrt{\text{Var}(Z_{u,k}(t))}} \geq m(u) \left(1 + (1 - \epsilon) \frac{B}{2A} \left(|k|^* \frac{\Delta(u)}{u} S\right)^2\right) := m_{k,\epsilon}(u)$$

with  $|k|^* = \min(|k|, |k+1|, |k+2|)$  as  $u$  sufficiently large. Thus for  $0 < \epsilon < 1$  and  $u$  sufficiently large

$$\Sigma_1^-(u) \leq \sum_{k=-N(u)-1}^{N(u)+1} \mathbb{P}(\Gamma(x, S; \bar{Z}_{u,k}) > m_{k,\epsilon}(u)).$$

Using that

$$\lim_{u \rightarrow \infty} \sup_{|k| \leq N(u)+1} \left| \frac{2\sigma^2(ut^*)}{\sigma^2(\Delta(u))(m_{k,\epsilon}(u))^2} - 1 \right| = 0,$$

by Lemma 3.2

$$(14) \quad \lim_{u \rightarrow \infty} \sup_{|k| \leq N(u)+1} \sup_{s \neq t, s, t \in [0, 2S]} \left| m_{k,\epsilon}^2(u) \frac{1 - \text{Cor}(Z_{u,k}(s), Z_{u,k}(t))}{\frac{\sigma^2(\Delta(u)|s-t|}{\sigma^2(\Delta(u))}} - 1 \right| = 0.$$

Hence in light of Proposition 2.3 in [6],

$$\lim_{u \rightarrow \infty} \sup_{|k| \leq N(u)+1} \left| \frac{\mathbb{P}(\Gamma(x, S; \bar{Z}_{u,k}(t)) > m_{k,\epsilon}(u))}{\Psi(m_{k,\epsilon}(u))} - \mathcal{H}_\eta^\Gamma(x, S) \right| = 0,$$

where  $\eta$  is defined in (7). Furthermore,

$$\begin{aligned} \Sigma_1^-(u) &\leq \sum_{k=-N(u)-1}^{N(u)+1} \mathcal{H}_\eta^\Gamma(x, S) \Psi(m_{k,\epsilon}(u)) \\ &\sim \mathcal{H}_\eta^\Gamma(x, S) \Psi(m(u)) \sum_{k=-N(u)-1}^{N(u)+1} e^{-(1-\epsilon)\frac{B}{2A}m^2(u)(|k|^* \frac{\Delta(u)}{u} S)^2} \\ &\sim \frac{\mathcal{H}_\eta^\Gamma(x, S)}{S} \frac{(1-\epsilon)^{-1/2} \left(\frac{B}{2A}\right)^{-1/2} u}{m(u)\Delta(u)} \Psi(m(u)) \int_{-\infty}^{\infty} e^{-t^2} dt \\ (15) \quad &\sim \frac{\mathcal{H}_\eta^\Gamma(x, S)}{S} \Theta(u), \quad u \rightarrow \infty, \epsilon \rightarrow 0, \end{aligned}$$

where

$$(16) \quad \Theta(u) := \sqrt{\frac{2A\pi}{B}} \frac{u}{m(u)\Delta(u)} \Psi(m(u)).$$

Similarly,

$$(17) \quad \Sigma_1^+(u) \geq \frac{\mathcal{H}_\eta^\Gamma(x, S)}{S} \Theta(u) (1 + o(1)), \quad u \rightarrow \infty.$$

Upper bound for  $\Sigma \Sigma_i(u)$ ,  $i = 1, 2$ . Similarly as in (14), Lemma 3.2 gives that

$$\lim_{u \rightarrow \infty} \sup_{s \neq t, s, t \in E_2(u)} \left| m^2(u) \frac{1 - \text{Cor}(Z_u(s), Z_u(t))}{\frac{\sigma^2(\Delta(u)|s-t|}{\sigma^2(\Delta(u))}} - 1 \right| = 0,$$

By Corollary 3.2 in [6], there exists  $\mathcal{C}, \mathcal{C}_1 > 0$  such that for all  $|k|, |l| \leq N(u) + 1, l \geq k + 2$ ,

$$\mathbb{P} \left( \sup_{t \in [kS, (k+1)S]} \bar{Z}_u(t) > m_{k,\epsilon}(u), \sup_{t \in [lS, (l+1)S]} \bar{Z}_u(t) > m_{l,\epsilon}(u) \right) \leq \mathcal{C} S^2 e^{-\mathcal{C}_1 |k-l|^\gamma S^\gamma} \Psi(\hat{m}_{u,k,l}),$$

with  $\gamma = \min(\alpha_0, \alpha_\infty)$  and  $\hat{m}_{k,l}(u) = \min(m_{k,\epsilon}(u), m_{l,\epsilon}(u))$ . Consequently, with aid of (15),

$$\begin{aligned} \Sigma \Sigma_2(u) &\leq \sum_{k=-N(u)-1}^{N(u)+1} \sum_{l=k+2}^{N(u)+1} \mathcal{C} S^2 e^{-\mathcal{C}_1 |k-l|^\gamma S^\gamma} \Psi(\hat{m}_{k,l}(u)) \\ &\leq \sum_{k=-N(u)-1}^{N(u)+1} \sum_{l=k+2}^{N(u)+1} \mathcal{C} S^2 e^{-\mathcal{C}_1 |k-l|^\gamma S^\gamma} (\Psi(m_{k,\epsilon}(u)) + \Psi(m_{l,\epsilon}(u))) \\ &\leq 2 \sum_{k=-N(u)-1}^{N(u)+1} \Psi(m_{k,\epsilon}(u)) \sum_{|k-l| \geq 1} \mathcal{C} S^2 e^{-\mathcal{C}_1 |k-l|^\gamma S^\gamma} \\ &\leq 2 \sum_{k=-N(u)-1}^{N(u)+1} \Psi(m_{k,\epsilon}(u)) \mathcal{C} S^2 e^{-\mathcal{Q} S^\gamma} \\ (18) \quad &\leq \mathcal{Q} S e^{-\mathcal{Q} S^\gamma} \Theta(u), \quad u \rightarrow \infty. \end{aligned}$$



Thus, again by Proposition 2.3 in [6], taking into account (14) and noting that  $\Gamma(0, S; f) = \sup_{t \in [0, S]} f(t)$ , we have

$$\lim_{u \rightarrow \infty} \sup_{|k| \leq N(u)+1} \left| \frac{\mathbb{P} \left( \sup_{t \in [0, S]} \overline{Z_{u,k}}(t) > \hat{m}_{k,k+1}(u) \right)}{\Psi(\hat{m}_{k,k+1}(u))} - \mathcal{H}_\eta([0, S]) \right| = 0.$$

Hence,

$$\begin{aligned} \hat{\Sigma}(u) &:= \sum_{k=-N(u)-1}^{N(u)+1} \mathbb{P} \left( \sup_{t \in [kS, (k+1)S]} Z_u(t) > m(u), \sup_{t \in [(k+1)S, (k+2)S]} Z_u(t) > m(u) \right) \\ &= \sum_{k=-N(u)-1}^{N(u)+1} \left( \mathbb{P} \left( \sup_{t \in [0, S]} \overline{Z_{u,k}}(t) > \hat{m}_{k,k+1}(u) \right) + \mathbb{P} \left( \sup_{t \in [0, S]} \overline{Z_{u,k+1}}(t) > \hat{m}_{k,k+1}(u) \right) \right) \\ &\quad - \sum_{k=-N(u)-1}^{N(u)+1} \mathbb{P} \left( \sup_{t \in [0, 2S]} \overline{Z_{u,k}}(t) > \hat{m}_{k,k+1}(u) \right) \\ &\leq \left( 2 \frac{\mathcal{H}_\eta([0, S])}{S} - \frac{\mathcal{H}_\eta([0, 2S])}{S} \right) \Theta(u), \quad u \rightarrow \infty, \end{aligned}$$

which together with (18) and the fact that

$$\Sigma \Sigma_1(u) = \Sigma \Sigma_2(u) + \hat{\Sigma}(u)$$

leads to

$$(19) \quad \Sigma \Sigma_1(u) \leq \left( 2 \frac{\mathcal{H}_\eta([0, S])}{S} - \frac{\mathcal{H}_\eta([0, 2S])}{S} + \mathbb{Q} S e^{-\mathbb{Q} S^\gamma} \right) \Theta(u), \quad u \rightarrow \infty.$$

Combination of (13), (15), (17)-(19) yields

$$(20) \quad \begin{aligned} \liminf_{u \rightarrow \infty} \frac{\pi_1(u)}{\Theta(u)} &\geq \frac{\mathcal{H}_\eta^\Gamma(x, S)}{S} - 2 \frac{\mathcal{H}_\eta([0, S])}{S} + \frac{\mathcal{H}_\eta([0, 2S])}{S} - \mathbb{Q} S e^{-\mathbb{Q} S^\gamma}, \\ \limsup_{u \rightarrow \infty} \frac{\pi_1(u)}{\Theta(u)} &\leq \frac{\mathcal{H}_\eta^\Gamma(x, S)}{S} + \mathbb{Q} S e^{-\mathbb{Q} S^\gamma}. \end{aligned}$$

Thus under the proviso that

$$(21) \quad \mathcal{H}_\eta^\Gamma(x) = \lim_{S \rightarrow \infty} \frac{\mathcal{H}_\eta^\Gamma(x, S)}{S} \in (0, \infty),$$

letting  $S \rightarrow \infty$  in (20) leads to

$$\lim_{u \rightarrow \infty} \frac{\pi_1(u)}{\Theta(u)} = \mathcal{H}_\eta^\Gamma(x) \in (0, \infty),$$

which combined with (9)-(11) establishes the claim.

Existence of  $\mathcal{H}_\eta^\Gamma(x)$ . In order to complete the proof, we are left with proving (21). By (20), we have

$$\liminf_{S \rightarrow \infty} \frac{\mathcal{H}_\eta^\Gamma(x, S)}{S} = \limsup_{S \rightarrow \infty} \frac{\mathcal{H}_\eta^\Gamma(x, S)}{S}.$$

By the fact that

$$\mathcal{H}_\eta^\Gamma(x, S) \leq \mathcal{H}_\eta([0, S]),$$

we have

$$\limsup_{S \rightarrow \infty} \frac{\mathcal{H}_\eta^\Gamma(x, S)}{S} \leq \mathcal{H}_\eta < \infty.$$

In order to prove positivity of  $\mathcal{H}_\eta^\Gamma(x)$ , we follow the same argument as in the proof of Theorem D.2 in [27]. Replacing  $\pi_1(u)$  in (12) by

$$\mathbb{P} \left( \exists s \in E_2(u), t \in E_2(u) \cap \bigcup_{k \in \mathbb{Z}} [2kS, (2k+1)S], s-t \geq x, Z_u(t) > m(u), Z_u(s) > m(u) \right)$$

and following the same arguments as for  $\pi_1(u)$ , we derive that for sufficiently large  $\hat{S} > x$

$$\liminf_{S \rightarrow \infty} \frac{\mathcal{H}_\eta^\Gamma(x, S)}{S} \geq \frac{\mathcal{H}_\eta([0, \hat{S}])}{\hat{S}} - \mathbb{Q}\hat{S}e^{-\mathbb{Q}(\hat{S})^\gamma} > 0,$$

where the last inequality follows by the fact that  $\mathcal{H}_\eta([0, S])$  is increasing with respect to  $S$ . Hence,

$$(22) \quad \lim_{S \rightarrow \infty} \frac{\mathcal{H}_\eta^\Gamma(x, S)}{S} \in (0, \infty).$$

This completes the proof.  $\square$

**3.2. Finite-time horizon.** Let

$$\Delta_1(u) = \frac{\sqrt{2}\sigma^2(T)}{u + cT}, \quad \Delta_2(u) = \left( \frac{\sigma(T)}{u + cT} \right)^2.$$

Observe that for any  $x \geq 0$

$$(23) \quad \mathbb{P} \left( \frac{\mathcal{T}_T^*(u) - \tau_T^*(u)}{\Delta_i(u)} > x \right) = \frac{\mathbb{P}(\mathcal{T}_T(u) - \tau_T(u) > x\Delta_i(u), \tau_T(u) \leq T)}{\mathbb{P}(\tau_T(u) \leq T)}, \quad i = 1, 2.$$

In the following lemma we give exact asymptotics of  $\mathbb{P}(\tau_T(u) \leq T) = \mathbb{P}(\sup_{t \in [0, T]} X(t) - ct > u)$ , referring for the proof to Theorem 2.3 and Remark 2.4 in [7] (choose  $\gamma = 0$ ).

**Lemma 3.3.** *Let  $X(t), t \in [0, T]$  be a centered Gaussian process with continuous trajectories and stationary increments satisfying BI-BII.*

i) *If  $t = o(\sigma^2(t)), t \rightarrow 0$ , then*

$$\mathbb{P}(\tau_T(u) \leq T) \sim \mathcal{H}_{B_{\alpha_0}}^\Gamma(0) \frac{\sigma^3(T)}{\dot{\sigma}(T)} (u^2 \Delta_1(u))^{-1} \Psi \left( \frac{u + cT}{\sigma(T)} \right).$$

ii) *If  $\sigma^2(t) \sim at, t \rightarrow 0, a > 0$ , then for  $d = \frac{2\sigma(T)\dot{\sigma}(T)}{a}$ ,*

$$\mathbb{P}(\tau_T(u) \leq T) \sim \mathcal{P}_{B_{1/2}}^d(0) \Psi \left( \frac{u + cT}{\sigma(T)} \right).$$

iii) *If  $\sigma^2(t) = o(t), t \rightarrow 0$ , then*

$$\mathbb{P}(\tau_T(u) \leq T) \sim \Psi \left( \frac{u + cT}{\sigma(T)} \right).$$

Thus in order to derive the limit of (23) as  $u \rightarrow \infty$ , it suffices to find the asymptotics of

$$\mathbb{P}(\mathcal{T}_T(u) - \tau_T(u) > x\Delta_i(u), \tau_T(u) \leq T), \quad i = 1, 2, \quad u \rightarrow \infty.$$

Let

$$\tilde{\sigma}_u(t) = \frac{\sigma(t)}{u + ct} \frac{u + cT}{\sigma(T)}, \quad 0 \leq t \leq T.$$

**Lemma 3.4.** *Suppose that **BI**, **BII** hold. Then for  $u$  sufficiently large,  $\tilde{\sigma}_u(t)$  attains its maximum over  $[0, T]$  at  $T$ , and for any  $\delta_u > 0$  with  $\lim_{u \rightarrow \infty} \delta_u = 0$*

$$\lim_{u \rightarrow \infty} \sup_{t \in [T - \delta_u, T]} \left| \frac{1 - \tilde{\sigma}_u(t)}{|t - T|} - \frac{\dot{\sigma}(T)}{\sigma(T)} \right| = 0.$$

Moreover,

$$\lim_{u \rightarrow \infty} \sup_{s \neq t, s, t \in [T - \delta_u, T]} \left| \frac{1 - \text{Cor}(X(s), X(t))}{\frac{\sigma^2(|s-t|)}{2\sigma^2(T)}} - 1 \right| = 0.$$

**Proof of Theorem 2.3** By (23) and Lemma 3.3 it suffices to analyze asymptotics of

$$\begin{aligned} & \mathbb{P}(\mathcal{T}_T(u) - \tau_T(u) > x\Delta_j(u), \tau_T(u) \leq T) \\ &= \mathbb{P}(\mathcal{T}_T(u) - \tau_T(u) > x\Delta_j(u), \tau_T(u) \leq T, X(T) - cT < u) \\ & \quad + \mathbb{P}(\mathcal{T}_T(u) - \tau_T(u) > x\Delta_j(u), \tau_T(u) \leq T, X(T) - cT > u) \\ &= \mathbb{P}(\exists 0 \leq s, t \leq T, s - t \geq x\Delta_j(u), X(t) - ct > u, X(s) - cs > u, X(T) - cT < u) \\ & \quad + \mathbb{P}(\exists 0 \leq t \leq T - x\Delta_j(u), X(t) - ct > u, X(T) - cT > u), \end{aligned}$$

where  $j = 1$  for cases i), ii) and  $j = 2$  for iii).

Thus, for  $j = 1, 2$

$$(24) \quad \sum_{i=3}^4 \pi_i^j(u) \leq \mathbb{P}(\mathcal{T}_T(u) - \tau_T(u) > x\Delta_j(u), \tau_T(u) \leq T) \leq \sum_{i=3}^4 \pi_i^j(u) + \pi_5(u),$$

where

$$\begin{aligned} \pi_3^j(u) &= \mathbb{P}(\exists T - (\ln u/u)^2 \leq s, t \leq T, s - t \geq x\Delta_j(u), X(t) - ct > u, X(s) - cs > u, X(T) - cT < u), \\ \pi_4^j(u) &= \mathbb{P}(\exists T - (\ln u/u)^2 \leq t \leq T - x\Delta_j(u), X(t) - ct > u, X(T) - cT > u), \\ \pi_5(u) &= \mathbb{P}\left(\sup_{t \in [0, T - (\ln u/u)^2]} X(t) - ct > u\right). \end{aligned}$$

Upper bound of  $\pi_5(u)$ . Rewrite  $\pi_5(u)$  as

$$\pi_5(u) = \mathbb{P}\left(\sup_{t \in [0, T - (\ln u/u)^2]} \frac{X(t) u + cT}{u + ct} \frac{u + cT}{\sigma(T)} > \frac{u + cT}{\sigma(T)}\right)$$

By the stationarity of increments of  $X$  and **BII**, we have

$$\mathbb{E}\left\{\left(\frac{X(t) u + cT}{u + ct} \frac{u + cT}{\sigma(T)} - \frac{X(s) u + cT}{u + cs} \frac{u + cT}{\sigma(T)}\right)^2\right\} \leq \mathbb{Q}(|t - s| + \sigma^2(|t - s|)) \leq \mathbb{Q}|t - s|^{\min(\alpha_0, 1)}$$

for  $s, t \in [0, T]$ . Moreover, following Lemma 3.4, we have that for  $u$  sufficiently large

$$\sup_{t \in [0, T - (\ln u/u)^2]} \text{Var}\left(\frac{X(t) u + cT}{u + ct} \frac{u + cT}{\sigma(T)}\right) \leq 1 - \mathbb{Q}_1(\ln u/u)^2.$$

Consequently, by Piterbarg inequality (see, e.g., Theorem 8.1 in [27]), for  $u$  sufficiently large

$$(25) \quad \pi_5(u) \leq \mathbb{Q}_2 u^{2/\min(\alpha_0, 1)} \Psi\left(\frac{u + cT}{\sigma(T) \sqrt{1 - \mathbb{Q}_1(\ln u/u)^2}}\right).$$

Asymptotics of  $\pi_i^j(u)$ ,  $i = 3, 4, j = 1, 2$ . Let

$$X_u(t) = \frac{X(T-t)}{u+c(T-t)} \frac{u+cT}{\sigma(T)}, \quad 0 \leq t \leq T.$$

Then for  $j = 1, 2$

$$\begin{aligned} \pi_3^j(u) &= \mathbb{P} \left( \exists 0 \leq s, t \leq (\ln u/u)^2, t-s \geq x\Delta_j(u), X_u(t) > \frac{u+cT}{\sigma(T)}, X_u(s) > \frac{u+cT}{\sigma(T)}, X_u(0) < \frac{u+cT}{\sigma(T)} \right), \\ \pi_4^j(u) &= \mathbb{P} \left( \exists x\Delta_j(u) \leq t \leq (\ln u/u)^2, X_u(t) > \frac{u+cT}{\sigma(T)}, X_u(0) > \frac{u+cT}{\sigma(T)} \right). \end{aligned}$$

In order to derive the asymptotics of  $\pi_i^j(u)$ ,  $i = 3, 4, j = 1, 2$ , we distinguish three scenarios: i)  $t = o(\sigma^2(t))$ , ii)  $\sigma^2(t) \sim at$  and iii)  $\sigma^2(t) = o(t)$  as  $t \rightarrow 0$ .

◇ Case i)  $t = o(\sigma^2(t))$ . Clearly,

$$(26) \quad \pi_4^1(u) \leq \mathbb{P} \left( X_u(0) > \frac{u+cT}{\sigma(T)} \right) = \Psi \left( \frac{u+cT}{\sigma(T)} \right).$$

We are left with deriving the asymptotics of  $\pi_3^1(u)$ .

Asymptotics of  $\pi_3^1(u)$ . We note that

$$\hat{\pi}_3(u) - \mathbb{P} \left( X_u(0) > \frac{u+cT}{\sigma(T)} \right) \leq \pi_3^1(u) \leq \hat{\pi}_3(u),$$

where

$$\hat{\pi}_3(u) = \mathbb{P} \left( \exists 0 \leq s, t \leq (\ln u/u)^2, t-s \geq x\Delta_1(u), X_u(t) > \frac{u+cT}{\sigma(T)}, X_u(s) > \frac{u+cT}{\sigma(T)} \right).$$

For  $S > x/2$ , let

$$(27) \quad X_{u,k}(t) = X_u(\Delta_1(u)(kS+t)), \quad N_1(u) = \left\lfloor \frac{(\ln u)^2}{u^2 \Delta_1(u) S} \right\rfloor - 1.$$

Bonferroni inequality gives that

$$\Sigma_2^-(u) - \Sigma \Sigma_3(u) \leq \hat{\pi}_3(u) \leq \Sigma_2^+(u),$$

where

$$\Sigma \Sigma_3(u) = \sum_{k=0}^{N_1(u)+1} \sum_{l=k+1}^{N_1(u)+1} \mathbb{P} \left( \sup_{t \in [kS, (k+1)S]} X_u(\Delta_1(u)t) > \frac{u+cT}{\sigma(T)}, \sup_{t \in [lS, (l+1)S]} X_u(\Delta_1(u)t) > \frac{u+cT}{\sigma(T)} \right),$$

and

$$\begin{aligned} \Sigma_2^\pm(u) &= \sum_{k=0}^{N_1(u) \pm 1} \mathbb{P} \left( \exists s \in [0, S], t \in [x+s, 2S], X_{u,k}(t) > \frac{u+cT}{\sigma(T)}, X_{u,k}(s) > \frac{u+cT}{\sigma(T)} \right) \\ &= \sum_{k=0}^{N_1(u) \pm 1} \mathbb{P} \left( \sup_{t \in [0, S]} \min \left( X_{u,k}(t), \sup_{s \in [x+t, 2S]} X_{u,k}(s) \right) > \frac{u+cT}{\sigma(T)} \right) \\ &= \sum_{k=0}^{N_1(u) \pm 1} \mathbb{P} \left( \Gamma(x, S; X_{u,k}) > \frac{u+cT}{\sigma(T)} \right), \end{aligned}$$

with  $\Gamma$  being defined in (3). By Lemma 3.4, we have that for any  $0 < \epsilon < 1$

$$\frac{\frac{u+cT}{\sigma(T)}}{\sup_{t \in [0, 2S]} \sqrt{\text{Var}(X_{u,k}(t))}} > \frac{u+cT}{\sigma(T)} \left( 1 + (1-\epsilon) \frac{\dot{\sigma}(T)}{\sigma(T)} |k| \Delta_1(u) S \right) := m_{k,\epsilon,1}(u)$$

as  $u$  sufficiently large. This implies that

$$\begin{aligned} \mathbb{P} \left( \Gamma(x, S; X_{u,k}) > \frac{u+cT}{\sigma(T)} \right) &\leq \mathbb{P} \left( \Gamma(x, S; \bar{X}_{u,k}) > \frac{\frac{u+cT}{\sigma(T)}}{\sup_{t \in [0, 2S]} \sqrt{\text{Var}(X_{u,k}(t))}} \right) \\ &\leq \mathbb{P} \left( \Gamma(x, S; \bar{X}_{u,k}) > m_{k,\epsilon,1}(u) \right). \end{aligned}$$

Moreover, by Lemma 3.4 we have

$$(28) \quad \lim_{u \rightarrow \infty} \sup_{0 \leq k \leq N_1(u)+1} \sup_{s \neq t, s, t \in [0, 2S]} \left| m_{k,\epsilon,1}^2(u) \frac{1 - \text{Cor}(X_{u,k}(t), X_{u,k}(s))}{\frac{\sigma^2(\Delta_1(u)|t-s|)}{\sigma^2(\Delta_1(u))}} - 1 \right| = 0.$$

Thus, following Proposition 2.3 in [6], we have

$$(29) \quad \sup_{|k| \leq N(u)+1} \left| \frac{\mathbb{P}(\Gamma(x, S; \bar{X}_{u,k}(t)) > m_{k,\epsilon,1}(u))}{\Psi(m_{k,\epsilon,1}(u))} - \mathcal{H}_{B_{\alpha_0}}^\Gamma(x, S) \right| = 0,$$

which combined with (22) implies that

$$\begin{aligned} \Sigma_2^+(u) &\leq \sum_{k=0}^{N_1(u)+1} \mathcal{H}_{B_{\alpha_0}}^\Gamma(x, S) \Psi(m_{k,\epsilon,1}(u)) \\ &\leq \mathcal{H}_{B_{\alpha_0}}^\Gamma(x, S) \Psi \left( \frac{u+cT}{\sigma(T)} \right) \sum_{k=0}^{N_1(u)+1} e^{-(1-\epsilon) \frac{\dot{\sigma}(T)}{\sigma^3(T)} |k| u^2 \Delta_1(u) S} \\ &\leq \frac{\mathcal{H}_{B_{\alpha_0}}^\Gamma(x, S)}{S} \frac{\sigma^3(T)}{(1-\epsilon) \dot{\sigma}(T) u^2 \Delta_1(u)} \Psi \left( \frac{u+cT}{\sigma(T)} \right) \int_0^\infty e^{-t} dt \\ &\sim \mathcal{H}_{B_{\alpha_0}}^\Gamma(x) \Theta_1(u), \quad u \rightarrow \infty, S \rightarrow \infty, \epsilon \rightarrow 0, \end{aligned}$$

with

$$\Theta_1(u) = \frac{\sigma^3(T)}{\dot{\sigma}(T)} (u^2 \Delta_1(u))^{-1} \Psi \left( \frac{u+cT}{\sigma(T)} \right).$$

Analogously,

$$\Sigma_2^-(u) \geq \mathcal{H}_{B_{\alpha_0}}^\Gamma(x) \Theta_1(u) (1 + o(1)), \quad u \rightarrow \infty, S \rightarrow \infty.$$

Following similar arguments as in (18)-(19), substituting  $\eta$  by  $B_{\alpha_0}$  and  $\Theta(u)$  by  $\Theta_1(u)$  in (18)-(19), we derive that

$$\begin{aligned} \Sigma \Sigma_3(u) &\leq \left( 2 \frac{\mathcal{H}_{B_{\alpha_0}}([0, S])}{S} - \frac{\mathcal{H}_{B_{\alpha_0}}([0, 2S])}{S} + \mathbb{Q} S e^{-\mathbb{Q} S^\gamma} \right) \Theta_1(u) \\ &= o(\Theta_1(u)), \quad u \rightarrow \infty, S \rightarrow \infty. \end{aligned}$$

Therefore,

$$\hat{\pi}_3(u) \sim \mathcal{H}_{B_{\alpha_0}}^\Gamma(x) \Theta_1(u), \quad u \rightarrow \infty.$$

By the fact that

$$\hat{\pi}_3(u) - \mathbb{P} \left( X_u(0) > \frac{u+cT}{\sigma(T)} \right) \leq \pi_3^1(u) \leq \hat{\pi}_3(u),$$

we have

$$\pi_3^1(u) \sim \mathcal{H}_{B_{\alpha_0}}^\Gamma(x)\Theta_1(u), \quad u \rightarrow \infty,$$

which combined with (24)-(26) leads to

$$\mathbb{P}(\mathcal{T}_T(u) - \tau_T(u) > x\Delta_j(u), \tau_T(u) \leq T) \sim \mathcal{H}_{B_{\alpha_0}}^\Gamma(x)\Theta_1(u), \quad u \rightarrow \infty.$$

Inserting the above and i) in Lemma 3.3 to (23), we establish the claim.

◇ ii) Case  $\sigma^2(t) \sim at$ . In this case we choose  $\Delta_1(u) = \frac{1}{\sigma} \left( \frac{\sqrt{2}\sigma^2(T)}{u+cT} \right)$  as the scaling function.

Asymptotics of  $\pi_3^1(u)$ . Using notation introduced in (27), we have for  $S > x$ ,

$$\pi_6(u) \leq \pi_3^1(u) \leq \pi_6(u) + \Sigma_3(u),$$

where

$$\begin{aligned} \pi_6(u) &= \mathbb{P}\left(\exists 0 \leq s \leq s+x \leq t \leq S, X_{u,0}(t) > \frac{u+cT}{\sigma(T)}, X_{u,0}(s) > \frac{u+cT}{\sigma(T)}, X_{u,0}(0) < \frac{u+cT}{\sigma(T)}\right), \\ &= \mathbb{P}\left(\Gamma'(x, S; X_{u,0}) > \frac{u+cT}{\sigma(T)}, X_{u,0}(0) < \frac{u+cT}{\sigma(T)}\right), \end{aligned}$$

with

$$\Gamma'(x, S; f) := \sup_{t \in [x, S]} \min\left(f(t), \sup_{s \in [0, t-x]} f(s)\right),$$

and

$$(30) \quad \Sigma_3(u) = \sum_{k=1}^{N_1(u)+1} \mathbb{P}\left(\sup_{t \in [0, S]} X_{u,k}(t) > \frac{u+cT}{\sigma(T)}\right).$$

Asymptotics of  $\pi_6(u)$ . We begin with observation that

$$\pi_6(u) = \mathbb{P}\left(\Gamma'\left(x, S; \frac{\bar{X}_{u,0}(t)}{1 + (1/\sqrt{\text{Var}(X_{u,0}(t))) - 1}}\right) > \frac{u+cT}{\sigma(T)}, X_{u,0}(0) < \frac{u+cT}{\sigma(T)}\right).$$

We shall apply Lemma 3.5 from Appendix, for which we verify assumptions **D0-D2** (see Appendix). Note that **D0** holds straightforwardly. From (28) for  $k=0$ , we know that **D1** is satisfied. By Lemma 3.4, it follows that

$$(31) \quad \lim_{u \rightarrow \infty} \sup_{t \in [0, S]} \left| \left(\frac{u+cT}{\sigma(T)}\right)^2 \left(1/\sqrt{\text{Var}(X_{u,0}(t))} - 1\right) - \frac{2\sigma(T)\dot{\sigma}(T)}{a}t \right| = 0.$$

This implies that **D2** is satisfied with  $h(t) = \frac{2\sigma(T)\dot{\sigma}(T)}{a}t$ . Moreover,

$$(32) \quad \left\{X_{u,0}(0) < \frac{u+cT}{\sigma(T)}\right\} = \left\{X_{u,0}(0) = \frac{u+cT}{\sigma(T)} - \frac{w}{\frac{u+cT}{\sigma(T)}}, w \in D\right\},$$

with  $D = (0, \infty)$ . Thus

$$\pi_6(u) \sim \int_0^\infty e^w \mathbb{P}(\Gamma'(x, S; W) > w) dw \Psi\left(\frac{u+cT}{\sigma(T)}\right), \quad u \rightarrow \infty.$$

with

$$(33) \quad W(t) = \sqrt{2}B_{1/2}(t) - \left(1 + \frac{2\sigma(T)\dot{\sigma}(T)}{a}\right)t.$$

Upper bound for  $\Sigma_3(u)$ . Noting that  $\Gamma(0, S; f) = \sup_{t \in [0, S]} f(t)$ , by (29) we have

$$\begin{aligned}
\Sigma_3(u) &\leq \sum_{k=1}^{N_1(u)+1} \mathbb{P} \left( \sup_{t \in [0, S]} \bar{X}_{u,k}(t) > m_{k,\epsilon,1}(u) \right) \\
&\leq \sum_{k=1}^{N_1(u)+1} \mathcal{H}_{B_{1/2}}([0, S]) \Psi(m_{k,\epsilon,1}(u)) \\
&\leq \mathcal{H}_{B_{1/2}}([0, S]) \Psi \left( \frac{u + cT}{\sigma(T)} \right) \sum_{k=1}^{N_1(u)+1} e^{-(1-\epsilon) \frac{\dot{\sigma}(T)}{\sigma^3(T)} k u^2 \Delta_1(u) S} \\
&\leq \mathbb{Q}S \sum_{k=1}^{N_1(u)+1} e^{-\mathbb{Q}_1 k S} \Psi \left( \frac{u + cT}{\sigma(T)} \right) \\
(34) \quad &\leq \mathbb{Q}S e^{-\mathbb{Q}_2 S} \Psi \left( \frac{u + cT}{\sigma(T)} \right) = o \left( \Psi \left( \frac{u + cT}{\sigma(T)} \right) \right), \quad u \rightarrow \infty, S \rightarrow \infty.
\end{aligned}$$

Therefore,

$$(35) \quad \pi_3^1(u) \sim \int_0^\infty e^w \mathbb{P} \left( \sup_{t \in [x, \infty)} \min \left( \sup_{s \in [0, t-x]} \eta(s), \eta(t) \right) > w \right) dw \Psi \left( \frac{u + cT}{\sigma(T)} \right), \quad u \rightarrow \infty.$$

Asymptotics of  $\pi_4^1(u)$ . Observe that

$$(36) \quad \pi_7(u) \leq \pi_4^1(u) \leq \pi_7(u) + \Sigma_3(u),$$

where  $\Sigma_3(u)$  is given by (30) and

$$\begin{aligned}
\pi_7(u) &= \mathbb{P} \left( \exists x \leq t \leq S, X_{u,0}(t) > \frac{u + cT}{\sigma(T)}, X_{u,0}(0) > \frac{u + cT}{\sigma(T)} \right) \\
&= \mathbb{P} \left( \sup_{t \in [x, S]} X_{u,0}(t) > \frac{u + cT}{\sigma(T)}, X_{u,0}(0) > \frac{u + cT}{\sigma(T)} \right).
\end{aligned}$$

Note that

$$\left\{ X_{u,0}(0) > \frac{u + cT}{\sigma(T)} \right\} = \left\{ X_{u,0}(0) = \frac{u + cT}{\sigma(T)} - \frac{w}{\frac{u + cT}{\sigma(T)}}, w \in D \right\}$$

with  $D = (-\infty, 0)$ . By (28), (31) and (32), applying Remark 3.6 in Appendix, it follows that

$$\pi_7(u) \sim \int_{-\infty}^0 e^w \mathbb{P} \left( \sup_{t \in [x, S]} W(t) > w \right) dw \Psi \left( \frac{u + cT}{\sigma(T)} \right),$$

with  $W$  given in (33). Inserting the above asymptotics and (34) into (26) gives that

$$(37) \quad \pi_4^1(u) \sim \int_{-\infty}^0 e^w \mathbb{P} \left( \sup_{t \in [x, S]} W(t) > w \right) dw \Psi \left( \frac{u + cT}{\sigma(T)} \right), \quad u \rightarrow \infty.$$

Combination of (24), (25), (35) and (37) gives the asymptotics for  $\mathbb{P}(\mathcal{T}_T(u) - \tau_T(u) > x \Delta_1(u), \tau_T(u) \leq T)$ , which together with ii) in Lemma 3.3 establishes the claim.

◇ Case iii)  $\sigma^2(t) = o(t)$ . In this case we choose  $\Delta_2(u) = \left( \frac{\sigma(T)}{u + cT} \right)^2$  as the scaling function.

Asymptotics of  $\pi_4^2(u)$ . Observe that

$$\pi_8(u) \leq \pi_4^2(u) \leq \pi_8(u) + \Sigma_4(u),$$

where

$$\begin{aligned}
\pi_8(u) &= \mathbb{P} \left( \exists x \leq t \leq S, \widehat{X}_{u,0}(t) > \frac{u+cT}{\sigma(T)}, \widehat{X}_{u,0}(0) > \frac{u+cT}{\sigma(T)} \right) \\
&= \mathbb{P} \left( \sup_{t \in [x, S]} \widehat{X}_{u,0}(t) > \frac{u+cT}{\sigma(T)}, \widehat{X}_{u,0}(0) > \frac{u+cT}{\sigma(T)} \right) \\
&= \mathbb{P} \left( \sup_{t \in [x, S]} \frac{\widehat{X}_{u,0}(t)}{1 + 1/\sqrt{\text{Var}(\widehat{X}_{u,0}(t)) - 1}} > \frac{u+cT}{\sigma(T)}, \widehat{X}_{u,0}(0) > \frac{u+cT}{\sigma(T)} \right) \\
\Sigma_4(u) &= \sum_{k=1}^{N_2(u)+1} \mathbb{P} \left( \sup_{t \in [0, S]} \widehat{X}_{u,k}(t) > \frac{u+cT}{\sigma(T)} \right) \\
&\leq \sum_{k=1}^{N_2(u)+1} \mathbb{P} \left( \sup_{t \in [0, S]} \widehat{\widehat{X}}_{u,k}(t) > m_{k,\epsilon,2}(u) \right),
\end{aligned}$$

where

$$\widehat{X}_{u,k}(t) = X_u(\Delta_2(u)(kS+t)), \quad N_2(u) = \left\lceil \frac{(\ln u)^2}{u^2 \Delta_2(u) S} \right\rceil,$$

$$m_{k,\epsilon,2}(u) = \frac{u+cT}{\sigma(T)} \left( 1 + (1-\epsilon) \frac{\dot{\sigma}(T)}{\sigma(T)} |k| \Delta_2(u) S \right).$$

*Asymptotics of  $\pi_8(u)$ .* In order to apply Lemma 3.5 and Remark 3.6 in Appendix, we check **D0-D2**.

Note that **D0** hold straightforwardly. By Lemma 3.4, it follows that

$$\lim_{u \rightarrow \infty} \sup_{|k| \leq N_2(u)+1} \sup_{s \neq t, s, t \in [0, S]} \left| (m_{k,\epsilon,2}(u))^2 \frac{1 - \text{Cor}(\widehat{X}_{u,k}(s), \widehat{X}_{u,k}(t))}{\frac{\sigma^2(\Delta_2(u)|t-s|)}{\sigma^2(\Delta_1(u))}} - 1 \right| = 0$$

with

$$\lim_{u \rightarrow \infty} \frac{\Delta_2(u)}{\Delta_1(u)} = 0.$$

This implies that **D1** holds with  $\nu = 0$ . Lemma 3.4 indicates that

$$\lim_{u \rightarrow \infty} \sup_{t \in [0, S]} \left| \left( \frac{u+cT}{\sigma(T)} \right)^2 \left( 1/\sqrt{\text{Var}(\widehat{X}_{u,0}(t)) - 1} \right) - \frac{\dot{\sigma}(T)}{\sigma(T)} t \right| = 0.$$

This means that **D2** holds with  $h(t) = \frac{\dot{\sigma}(T)}{\sigma(T)} t$ . Moreover,

$$\left\{ \widehat{X}_{u,0}(0) > \frac{u+cT}{\sigma(T)} \right\} = \left\{ \widehat{X}_{u,0}(0) = \frac{u+cT}{\sigma(T)} - \frac{w}{\frac{u+cT}{\sigma(T)}}, w \in D \right\},$$

with  $D = (-\infty, 0)$ . Hence

$$\begin{aligned}
\pi_8(u) &\sim \int_{-\infty}^0 e^w \mathbb{P} \left( \sup_{t \in [x, S]} -\frac{\dot{\sigma}(T)}{\sigma(T)} t > w \right) dw \Psi \left( \frac{u+cT}{\sigma(T)} \right) \\
&= e^{-\frac{\dot{\sigma}(T)}{\sigma(T)} x} \Psi \left( \frac{u+cT}{\sigma(T)} \right).
\end{aligned}$$



Upper bound of  $\Sigma_4(u)$ . Similarly, for the summands in  $\Sigma_4(u)$  we can show that **D0-D2** hold with  $h(t) = \frac{\dot{\sigma}(T)}{\sigma(T)}t$  and  $\nu = 0$ . Thus by Remark 3.6 in Appendix, we have

$$\lim_{u \rightarrow \infty} \sup_{|k| \leq N_2(u)+1} \left| \frac{\mathbb{P} \left( \sup_{t \in [0, S]} \widehat{X}_{u, k}(t) > m_{k, \epsilon, 2}(u) \right)}{\Psi(m_{k, \epsilon, 2}(u))} - \int_{-\infty}^{\infty} e^w \mathbb{P} \left( \sup_{t \in [0, S]} -\frac{\dot{\sigma}(T)}{\sigma(T)}t > w \right) dw \right| = 0,$$

where

$$\int_{-\infty}^{\infty} e^w \mathbb{P} \left( \sup_{t \in [0, S]} -\frac{\dot{\sigma}(T)}{\sigma(T)}t > w \right) dw = 1.$$

Hence

$$\begin{aligned} \Sigma_4(u) &\leq \sum_{k=1}^{N_2(u)+1} \Psi(m_{k, \epsilon, 2}(u)) \\ &\leq \Psi \left( \frac{u + cT}{\sigma(T)} \right) \sum_{k=1}^{N_2(u)+1} e^{-\mathbb{Q}|k|\Delta_2(u)u^2 S} \\ &\leq \Psi \left( \frac{u + cT}{\sigma(T)} \right) \sum_{k=1}^{N_2(u)+1} e^{-\mathbb{Q}_1 S} \\ &\leq e^{-\mathbb{Q}_2 S} \Psi \left( \frac{u + cT}{\sigma(T)} \right) = o \left( \Psi \left( \frac{u + cT}{\sigma(T)} \right) \right), \quad u \rightarrow \infty, S \rightarrow \infty. \end{aligned}$$

Thus

$$\pi_4^2(u) \sim e^{-\frac{\dot{\sigma}(T)}{\sigma(T)}x} \Psi \left( \frac{u + cT}{\sigma(T)} \right), \quad u \rightarrow \infty.$$

Upper bound of  $\pi_3^2(u)$ . It follows that

$$\begin{aligned} \pi_3^2(u) &\leq \mathbb{P} \left( \tau_T(u) \leq T, X_u(0) < \frac{u + cT}{\sigma(T)} \right) \\ &= \mathbb{P}(\tau_T(u) \leq T) - \mathbb{P} \left( \tau_T(u) \leq T, X_u(0) > \frac{u + cT}{\sigma(T)} \right) \\ &= \mathbb{P}(\tau_T(u) \leq T) - \mathbb{P} \left( X_u(0) > \frac{u + cT}{\sigma(T)} \right) \end{aligned}$$

Applying iii) in Lemma 3.3, we have

$$\pi_3^2(u) = o \left( \Psi \left( \frac{u + cT}{\sigma(T)} \right) \right), \quad u \rightarrow \infty.$$

Recalling (24)-(25), we conclude that

$$\mathbb{P}(\mathcal{T}_T(u) - \tau_T(u) > x\Delta_2(u), \tau_T(u) \leq T) \sim e^{-\frac{\dot{\sigma}(T)}{\sigma(T)}x} \Psi \left( \frac{u + cT}{\sigma(T)} \right), \quad u \rightarrow \infty,$$

which combined with (23) and iii) in Lemma 3.3 leads, for any  $x > 0$ , to

$$\lim_{u \rightarrow \infty} \mathbb{P} \left( \frac{\mathcal{T}_T^*(u) - \tau_T^*(u)}{\Delta_2(u)} > x \right) = e^{-\frac{\dot{\sigma}(T)}{\sigma(T)}x}.$$

This completes the proof.  $\square$

## APPENDIX

In this section we give a variant of Theorem 2.1 in [6]. Let  $\xi_{u,\tau_u}$  be a family of Gaussian random fields given by

$$(38) \quad \xi_{u,\tau_u}(t) = \frac{Z_{u,\tau_u}(t)}{1 + h_{u,\tau_u}(t)}, \quad t \in [0, S], \tau_u \in K_u,$$

where  $Z_{u,\tau_u}$  is a family of centered Gaussian random fields with continuous trajectories and unit variance,  $h_{u,\tau_u} \in C([0, S])$ ,  $S > 0$  and  $K_u$  is a set of index. We investigate the asymptotics of

$$\mathbb{P}(\Gamma'(x, S; \xi_{u,\tau_u}) > g_{u,\tau_u}, D_{u,\tau_u})$$

as  $u \rightarrow \infty$  where  $g_{u,\tau_u}$  is a series of positive functions of  $u$ ,

$$(39) \quad D_{u,\tau_u} = \{Z_{u,\tau_u}(0) = g_{u,\tau_u} - \frac{w}{g_{u,\tau_u}}, w \in D\}$$

with  $D = (0, \infty), (-\infty, 0)$  or  $D = \mathbb{R}$  and  $\Gamma' : C([0, S]) \rightarrow \mathbb{R}$ ,  $0 \leq x \leq S$  is a real-valued continuous functional defined by

$$(40) \quad \Gamma'(x, S; f) = \sup_{t \in [x, S]} \min \left( f(t), \sup_{s \in [0, t-x]} f(s) \right), \quad f \in C([0, S]).$$

In order to avoid trivialities, we assume that

$$\lim_{u \rightarrow \infty} \mathbb{P}(\Gamma'(x, S; \xi_{u,\tau_u}) > g_{u,\tau_u}, D_{u,\tau_u}) = 0$$

and observe that  $\mathbb{P}(\Gamma'(x, S; \xi_{u,\tau_u}) > g_{u,\tau_u}, D_{u,\tau_u}) = \mathbb{P}(\Gamma'(x, S; \xi_{u,\tau_u}) > g_{u,\tau_u})$  if  $D = \mathbb{R}$ .

As in [6] (see Theorem 2.1), we impose the following assumptions:

**D0:**  $\lim_{u \rightarrow \infty} \inf_{\tau_u \in K_u} g_{u,\tau_u} = \infty$ .

**D1:** There exist  $\rho(t)$ , regularly varying function at 0 with index  $2\alpha_0 \in (0, 2]$  and  $b_i(u) > 0, i = 1, 2$  satisfying  $\lim_{u \rightarrow \infty} b_i(u) = 0, i = 1, 2$  and  $\lim_{u \rightarrow \infty} \frac{b_1(u)}{b_2(u)} = \nu \in [0, \infty)$  such that

$$\lim_{u \rightarrow \infty} \sup_{\tau_u \in K_u} \sup_{s, t \in [0, S], s \neq t} \left| (g_{u,\tau_u})^2 \frac{1 - \text{Corr}(Z_{u,\tau_u}(t), Z_{u,\tau_u}(s))}{\frac{\rho(b_1(u)|t-s|)}{\rho(b_2(u))}} - 1 \right| = 0.$$

**D2:** There exists  $h \in C([0, S])$  such that

$$\lim_{u \rightarrow \infty} \sup_{\tau_u \in K_u} \sup_{t \in [0, S]} |(g_{u,\tau_u})^2 h_{u,\tau_u}(t) - h(t)| = 0.$$

**Lemma 3.5.** *Let  $\xi_{u,\tau_u}$  be defined as in (38) and  $\Gamma'$  be defined in (40). Assume that **D0-D2** are satisfied. Then, for  $D_{u,\tau_u}$  defined in (39) with  $D = (0, \infty), (-\infty, 0)$  or  $D = \mathbb{R}$ ,*

$$\lim_{u \rightarrow \infty} \sup_{\tau_u \in K_u} \left| \frac{\mathbb{P}(\Gamma'(x, S; \xi_{u,\tau_u}) > g_{u,\tau_u}, D_{u,\tau_u})}{\Psi(g_{u,\tau_u})} - \int_D e^w \mathbb{P}(\Gamma'(x, S; \nu^{\alpha_0} B_{\alpha_0}(t) - \nu^{2\alpha_0} |t|^{2\alpha_0} - h(t)) > w) dw \right| = 0.$$

**Proof.** Conditioning on the event that  $Z_{u,\tau_u}(0) = g_{u,\tau_u} - \frac{w}{g_{u,\tau_u}}$  and noting that  $D_{u,\tau_u} = \{Z_{u,\tau_u}(0) = g_{u,\tau_u} - \frac{w}{g_{u,\tau_u}}, w \in D\}$ , we have

$$\begin{aligned} & \mathbb{P}(\Gamma'(x, S; \xi_{u,\tau_u}) > g_{u,\tau_u}, D_{u,\tau_u}) \\ &= \frac{1}{\sqrt{2\pi}g_{u,\tau_u}} \int_{\mathbb{R}} e^{-\frac{(g_{u,\tau_u} - \frac{w}{g_{u,\tau_u}})^2}{2}} \mathbb{P}\left(\Gamma'(x, S; \xi_{u,\tau_u}) > g_{u,\tau_u}, D_{u,\tau_u} \mid Z_{u,\tau_u}(0) = g_{u,\tau_u} - \frac{w}{g_{u,\tau_u}}\right) dw \end{aligned}$$

$$= \frac{e^{-\frac{(g_{u,\tau_u})^2}{2}}}{\sqrt{2\pi}g_{u,\tau_u}} \int_D e^{w-\frac{w^2}{2(g_{u,\tau_u})^2}} \mathbb{P}\left(\Gamma'(x, S; \xi_{u,\tau_u}) > g_{u,\tau_u} \middle| Z_{u,\tau_u}(0) = g_{u,\tau_u} - \frac{w}{g_{u,\tau_u}}\right) dw.$$

Using the same procedure as in the proof of Theorem 2.1 in [6], we can show that

$$\int_D e^{w-\frac{w^2}{2(g_{u,\tau_u})^2}} \mathbb{P}\left(\Gamma'(x, S; \xi_{u,\tau_u}) > g_{u,\tau_u} \middle| Z_{u,\tau_u}(0) = g_{u,\tau_u} - \frac{w}{g_{u,\tau_u}}\right) dw$$

uniformly converges to

$$\int_D e^w \mathbb{P}\left(\Gamma'(x, S; \nu^{\alpha_0} B_{\alpha_0}(t) - \nu^{2\alpha_0} |t|^{2\alpha_0} - h(t)) > w\right) dw$$

with respect to  $\tau_u \in K_u$ . This completes the proof.  $\square$

**Remark 3.6.** Lemma 3.5 also holds if we substitute  $\Gamma'$  by  $\sup_{t \in [x, S]} f(t)$ , with  $f \in C[x, S]$ ,  $x \geq 0$  and  $D$ , a measurable subset of  $\mathbb{R}$  with positive Lebesgue measure.

**Acknowledgement:** We thank Enkelejd Hashorva for discussions and comments that improved presentation of the results of this contribution. K. Dębicki was partially supported by NCN Grant No 2015/17/B/ST1/01102 (2016-2019) whereas P. Liu was supported by the Swiss National Science Foundation Grant 200021-175752/1.

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