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The relationship between cognitive decline and a genetic predictor of educational attainment

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3 Abstract

4 Genetic and environmental factors both make substantial contributions to the heterogeneity 5 in individuals' levels of cognitive ability. Many studies have examined the relationship 6 between educational attainment and cognitive performance and its rate of change. Yet there 7 remains a gap in knowledge regarding whether the effect of genetic predictors on individual 8 differences in cognition becomes more or less prominent over the life course. In this analysis 9 of over 5,000 older adults from the Health and Retirement Study (HRS) in the U.S., we measured the change in performance on global cognition, episodic memory, attention & 10 11 concentration, and mental status over 14 years. Growth curve models are used to evaluate the association between a polygenic risk score for education (education PGS) and cognitive 12 13 change. Using the most recent education PGS, we find that individuals with higher scores 14 perform better across all measures of cognition in later life. Education PGS is associated 15 with a faster decline in episodic memory in old age. The relationships are robust even after 16 controlling for phenotypic educational attainment, and are unlikely to be driven by mortality 17 bias. Future research should consider genetic effects when examining non-genetic factors in 18 cognitive decline. Our findings represent a need to understand the mechanisms between 19 genetic endowment of educational attainment and cognitive decline from a biological angle.

20 Keywords:

- 21 Cognitive decline, Educational Attainment, Fluid/Crystallised Intelligence, Genetic
- 22 predictor, Growth curve modelling; Polygenic risk score

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24 Cognitive competencies tend to decline with age. Interpersonal variability in age-related 25 cognitive decline is not fully understood: while some people experience substantial deterioration in cognitive function, others maintain better cognitive status despite the 26 presence of considerable brain deterioration (Stern, 2009). Cognitive decline threatens 27 independence and quality of life for older adults (Williams & Kemper, 2010). With an ageing 28 population, both in the U.S. and worldwide, cognitive decline is an emerging health and 29 social issue, especially since older individuals are increasingly taking additional 30 responsibility for financial and medical decisions. Understanding the predictors that 31 contribute to the variation in the trajectories of cognitive ageing has important biological and 32 public health implications. It may not only provide insight into the deterioration of cognitive 33 34 function, but also enable us to identify individuals at high risk of rapid decline and the development of personalised strategies for prevention of cognitive-skill related comorbidities. 35

36 Genetic, socioeconomic and behavioural risk factors all make substantial contributions to the heterogeneity in individuals' level of cognitive ability. Twin and family-based studies 37 indicate that at least moderate proportion of the differences in most domains of cognitive 38 39 ability is associated with genetic factors (Bouchard & McGue, 2003; Rietveld et al., 2014). Social scientists have shown that the relationship between education and cognition is in part 40 due to the causal effect of schooling. This relationship can also be due to genetic 41 42 confounding. Recent genome-wide association studies (GWAS) found that the genetic 43 components of general cognitive functions are about 20-30% heritable (Davies et al., 2016). Higher educational attainment may allow individuals to cope more effectively with age-44 45 related brain deterioration, and thus perform better on cognitive tasks in later life (Lenehan, Summers, Saunders, Summers, & Vickers, 2015; Rietveld et al., 2014; Scarmeas & Stern, 46 47 2004). Recent GWAS have discovered molecular genetic associations with education (Lee et

al., 2018) and general cognition (Davies et al., 2018). The polygenic score of education
constructed by Lee et al. (2018) explains 11-13% of the variance of educational attainment
and 7-10% of the variance of cognitive performance, suggesting that the phenotypes have
shared genetic basis (Marioni et al., 2014; Okbay, Beauchamp, Fontana, Lee, & Pers, 2016;
Rietveld et al., 2014). These findings suggest that common genetic effects may account for
some of the observed association between education and cognitive ability.

Genetic variants that promote educational attainment – those that influence brain 54 development and neuron-to-neuron communication, for example – may have an effect on 55 cognitive functioning throughout the life course (Lee et al., 2018). The magnitude of their 56 effects may also change with age. For example, previous research attempting to examine the 57 58 importance of genetic risks across the life course has shown that ageing magnifies genetic 59 effects on cognitive ability (Laukka et al., 2013; Li et al., 2013; Papenberg et al., 2014; Papenberg, Lindenberger, & Bäckman, 2015). The rationale is that the association between 60 brain resources and cognitive ability is nonlinear, and that genetic variation is more 61 influential on performance differences during normal ageing when the brain starts to lose 62 neurochemical and structural resources (Lindenberger et al., 2008). However, the majority of 63 previous studies suffers from shortcomings regarding research design and methodology, 64 including cross-sectional data sources (Li et al., 2013; Nagel, 2008), a small number of 65 assessed genetic variants (Bretsky, Guralnik, Launer, Albert, & Seeman, 2003; Schiepers et 66 67 al., 2012), and focus on a narrow period of the life course (Moorman, Carr, & Greenfield, 2018). Therefore, it remains largely unknown how affect the trajectory of cognitive abilities 68 69 and its rate of change among older adults.

The objective of this study is thus to investigate whether the polygenic score for educaion are
associated with later life cognitive functions and cognitive decline independently among

72 middle-aged and older adults in the United States. We measure cognition and its decline both 73 separately in the domains of episodic memory, attention and calculation, and mental status, and as an index measuring general cognition. We use polygenic scores constructed for the 74 75 Health and Retirement Study (HRS) that summarise an individual's cumulative genetic predictor to educational attainment. The polygenic scores for educational attainment 76 77 (hereafter, education PGS) are constructed by adding the effect-size-weighted risk alleles across the genome associated with education based on the third and most recent educational 78 79 attainment GWAS consortium paper by Lee et al. (2018), which used data from 1.1 million participants and identified 1,271 lead genetic variants. The education PGS correlats with 80 81 years of education ($\beta = 0.8$; se = 0.03) with a predictive power of 10% in our HRS sample 82 (see supplementary material for more details). This research tackles the following three research questions: 1) How are the education PGS associated with level of cognitive function? 83 2) How does the effect of education PGS on individual differences in cognition change with 84 age? 3) Does the relationship between genes, age, and cognition still hold after controlling for 85 other socioeconomic, behavioural and health factors? We use growth curve analysis across 86 87 the waves (1998-2014) of the HRS to gain a better understanding of how genes and education operate across the life course as people age. 88

89 Genetics Predictors of Educational Attainment and Cognitive Decline

Cognitive ability varies among individuals across the life-span. Moreover, the within-person sub-dimensions of cognitive decline at different rates: verbal, numerical and knowledgebased abilities remain relatively stable in late life, while other mental abilities such as memory and processing speed start to deteriorate from middle age or even earlier and at a faster rate (Mustafa et al., 2012; Nisbett et al., 2012). Episodic memory – the ability to encode and retrieve personally experienced events that occurred at a specific place and time

96 (Gabrieli, 1998) – is a type of fluid intelligence that involves the ability to think and reason 97 abstractly. Evidence suggests that episodic memory is independent of pre-existing knowledge, learning and education, and is relatively more sensitive to genetic variability 98 (Smith et al., 2018), for example, the Apolipoprotein E (APOE). On the other hand, mental 99 100 status, attention and calculation are types of crystallised intelligence, which is formed through accumulating knowledge and experience. As people age and gain new knowledge 101 and understanding, crystallised intelligence tends to increase first and decline more slowly 102 (Salthouse, 2012). As a channel to gain knowledge and skills, education is expected to have a 103 104 substantial effect on crystallised abilities.

105 Wedow et al. (2018) described two pathways (Figure 1) through which genes could influence 106 social and health outcomes. One pathway for the connection between education PGS and cognitive functions is *biological pleiotropy* (Pickrell et al., 2016). That is, genes contribute to 107 both educational attainment and cognition independently due to underlying biological and 108 latent genetic mechanisms. For example, a person's genetics may influence brain 109 development to affect non-cognitive self-control, interpersonal skills, preferences and 110 111 behaviours leading to differences in both educational attainment and cognition (Belsky et al., 112 2016; Okbay et al., 2016). In addition, a person's education-increasing genotypes may be 113 associated with parental education-increasing genotypes, which can in turn select the 114 individuals into socially advantaged families that promote both cognitive development and 115 educational attainment (Belsky et al., 2016, 2018; Kong et al., 2018; Lee et al., 2018). This explanation raises the issue that any observed association between educational attainment and 116 117 cognition may be spurious due to the omitted genetic variable bias.

118 The second pathway for the association between genetic predictors for education and 119 cognitive status is *mediated pleiotropy*. Well-established evidence suggests that educational

120 level in early life affects the level of cognitive performance in later life (Gatz et al., 2001; 121 Tucker-Drob, Johnson, & Jones, 2009; Van Dijk, Van Gerven, Van Boxtel, Van der Elst, & Jolles, 2008; Zahodne et al., 2011). The potential mechanisms are improved cultural 122 123 competence and reasoning skills, a more effective use of brain function and cognitive processing, and a healthier occupational environment and lifestyle (Andel, Vigen, Mack, 124 125 Clark, & Gatz, 2006; Chen, Anthony, & Crum, 1999; Kramer, Bherer, Colcombe, Dong, & Greenough, 2004). According to this explanation, genes that are causally associated with 126 127 education affect cognitive performance through the mediated path.

The first research question is therefore whether there is an association between education 128 PGS and the level of cognitive ability? Evaluating this relationship could aid our 129 130 understanding of the association between educational attainment and cognitive status - to 131 what extent cognition is influenced by educational attainment via biological mechanisms and 132 unobservable confounders related to environmental factors. Drawing from recent advances in GWAS and the theoretical relationship between education and cognitive performance, we 133 hypothesise that education PGS is positively related to the level of cognitive ability in old 134 135 age, independent of the phenotypic educational attainment (Hypothesis 1).

The relationship between genetics and cognitive rate of change is less straightforward and 136 more domain-specific. Since crystallised intelligence is more dependent on education, we 137 138 speculate that genetics are more likely to affect attention and concentration, and mental status 139 via mediated pleiotropy. Evidence on whether educational level influences the trajectory of age-related cognitive decline is inconsistent. These inconsistent findings may be due to 140 141 methodological differences, such as sample characteristics, analytic strategies, type of cognitive measures and decline, or selection and confounding (Foverskov et al., 2018; 142 143 Gottesman et al., 2014).

144 Some earlier studies linking education with cognitive change in old age find that lower levels of education are associated with a faster decline in verbal fluency, mental status and general 145 cognition (e.g. Albert et al., 1995; Jacqmin-Gadda, Fabrigoule, Commenges, & Dartigues, 146 1997; Lyketsos, Chen, & Anthony, 1999). These studies posit that individuals with a higher 147 level of education use brain networks or cognitive paradigms more efficiently or flexibly, and 148 would exhibit a smaller decline in cognitive function relative to those with a lower level of 149 education (Salthouse, Atkinson, & Berish, 2003). More recent studies cast doubts on whether 150 151 rates of cognitive decline vary by education in later life. Many suggest that higher levels of education do not attenuate the rate of decline in episodic memory, working memory, 152 153 processing speed and verbal fluency (Glymour, Tzourio, & Dufouil, 2012; Gottesman et al., 154 2014; Karlamangla et al., 2009; Zahodne et al., 2011). Others report that higher education is associated with faster cognitive decline in attention & concentration. (Gottesman et al., 155 156 2014; Zahodne, Stern, & Manly, 2015). A potential explanation for the lack of positive association between education and rate of cognitive decline is that education raises baseline 157 cognitive performance, which increases the time needed to decline to the pathological 158 159 threshold. People with higher level of education thus decline at a similar rate to their lowereducated counterparts, or even a faster rate if they rely on specific cognitive domains to 160 161 compensate for declines in other cognitive domains. In summary, recent evidence shows no association on the phenotypic educational attainment and cognitive decline. Since declines in 162 163 crystallised domains are less sensitive to the educational protective effect, we extend the 164 phenotypic perspective to genetic inquiry and hypothesise that education PGS are not associated with the rate of cognitive change in crystallised domains (Hypothesis 2). 165

166 From a biological perspective, studies demonstrate magnified genetic influence on different167 types of cognition and the rate of cognitive decline during normal ageing (Tucker-Drob,

168 Reynolds, Finkel, & Pedersen, 2014). Meta-analysis suggests increased heritability for 169 episodic memory, working memory and spatial ability from early to late adulthood (Reynolds & Finkel, 2015). The resource-modulation hypothesis proposed by Lindenberger et al., 170 (2008) hypothesises that losses of structural and neurochemical brain resources in non-171 pathological ageing moderate the effects of common genetic variations on cognitive 172 performance. The hypothesis assumes a non-linear function linking brain resources to 173 cognitive abilities, and differences in genetic predictor exert magnifying effects on cognitive 174 functions as brain resources reduce from high to medium levels. Given that episodic memory 175 may be closer to the molecular effects of a gene than cognitive reserve such as education, the 176 177 rate of change in episodic memory is expected to be more sensitive to genetic effects 178 (Papenberg, Lindenberger, & Bäckman, 2015; Rasch, Papassotiropoulos, & de Quervain, 2010). Older adults, therefore, may benefit more from positive genetic endowment relative to 179 180 their younger counterparts. Therefore, we hypothesise that a higher level of education PGS 181 may be associated with a lower rate of decline in episodic memory (Hypothesis 3).

Previous research has found evidence that supports the resource-modulation hypothesis. For 182 183 example, APOE polymorphism is involved in lipid homeostasis and injury repair in the brain 184 (Papenberg et al., 2015): carrying the ɛ4 allele is a strong risk factor for accelerated cognitive decline in ageing (Filippini et al., 2011; Liu et al., 2010; Zhang & Pierce, 2014). However, 185 186 the literature to date suffers from a few limitations. First, the majority of past research 187 focuses only on one or a handful of genetic variants such as the aforementioned APOE (Bretsky et al., 2003; Schiepers et al., 2012), brain-derived neurotrophic factor (Ghisletta et 188 189 al., 2014), catechol-O-methyltransferas (Papenberg et al., 2013) and kidney- and brain-190 expressed protein (Muse et al., 2014). These studies are controversial as they tend to produce results that are rarely replicable due to their lack of power to detect plausible effects 191

(Benjamin et al., 2012; Chabris et al., 2012). Second, a large number of studies adopt a crosssectional design (Li et al., 2013; Nagel, 2008); longitudinal studies are rare but necessary to confirm the patterns observed in the cross-sectional data (Papenberg et al., 2015). Third, recent studies using the polygenic scores from GWAS studies and longitudinal data sets tend to focus on cognitive development in young age (Moorman et al., 2018), and at a narrow period of the life course (Ritchie et al., 2019).

In summary, our research studies how genetic variants influence trajectories of cognitive performance across the later lifespan. We overcome the aforementioned limitations by measuring genetic predictors for education using the polygenic score method among over 7,000 individuals aged 50 and above and tracked over 16 years.

202 Data and Methods

203 Data

204 The Health and Retirement Study (HRS) began in 1992 and is a biennial, longitudinal survey of a nationally-representative sample of individuals and their spouses aged 50 and above. In 205 2006 and 2008, the HRS collected genetic (saliva) samples from approximately 84% of 206 participants undergoing face-to-face interviews (12,507 individuals). These DNA samples 207 208 were genotyped for about 2 million SNPs. This study exploits the longitudinal nature of the 209 HRS to explore cognitive performance trajectories among older adults in the U.S. We use 210 eight waves of HRS data (from 1998 to 2012). Pre-processed datasets included the user-211 friendly RAND HRS data files (version P) and 1998-2012 HRS Core Files.

212 Sample

213 During the period 1998-2012, 8,652 respondents were genotyped. Growth curve models 214 typically require at least three waves of repeated measures for each individual (Curran, 215 Obeidat, & Losardo, 2010), 2,699 (31.2%) respondents whose cognitive performance was 216 measured fewer than three times were removed. Since this study only focuses on age-related 217 cognitive decline, we have differentiated normal cognitive functioning from impaired 218 functioning. A composite score measuring memory and mental status have been constructed 219 (ranging from 0 - 27). Respondents (n=36) with a score of less than seven exhibited signs of dementia (Crimmins, Kim, Langa, & Weir, 2011) and were removed. Finally, for the main 220 221 analysis, only individuals from European and non-Hispanic backgrounds were included. The 222 5,859 remaining respondents had at least three cognitive interviews: 25% had four or fewer 223 interviews, 50% had six or more interviews, providing 34,184 person-wave observations.

224 Dependent variables - cognitive measures

In the HRS, assessment of cognitive function is based on a reduced version of the telephone interview for the assessment of cognitive status (Desmond, Tatemichi, & Hanzawa, 1994), which was derived from the Mini-Mental State Exam (MMSE) (Folstein, Folstein, & McHugh, 1975). The assessment has been validated for use as a screening instrument for cognitive performance. The same cognitive tests were administered during all the included waves of data collection and were used to construct cognitive trajectories for individuals on each test (Herzog & Wallace, 1997).

Episodic memory (EM) was measured by immediate and delayed word recall. Respondents were read a list of ten common words (e.g. hotel, sky, water) and were then asked to recall as

many of them as possible both immediately after the list was read and also several minutes
later. The score records the total number of words the respondent correctly recalled at each
instance and ranges from 0 to 20.

Attention & calculation (A&C) was assessed with the serial 7s subtraction test. The respondents were asked to subtract 7 from 100 and continue subtracting 7 from each subsequent number for a total of five trials. The scores record the correct number of trials (ranging from 0-5). The serial 7s subtraction test assessed mixed abilities of attention, calculation and working memory that maintains and manipulates information using shortterm memory.

Mental status (MS) was assessed by naming the date, month, year and day of the week (ranging from 0-4), backwards counting from 20 (0-2), object naming (0-2), and naming the current president and vice president of the U.S (0-2).

Global cognition (GC) is a summary measure of the cognitive domains mentioned above (ranging from 0-35). To provide comparability across all measurements, we rescaled individual and global cognitive variables into a corrected percentage score – based on division by the maximum score and multiplication by 100.

The HRS includes other additional cognitive measures such as Wechsler Adult Intelligence Scale similarities, numeracy, quantitative reasoning and verbal fluency modules. We chose not to include them in our analyses as they were either asked of a small group of respondents or only added to the survey waves recently (Fisher, Hassan, Faul, Rodgers, & Weir, 2017).

254 Independent variable – education PGS

The education PGS is based on the most recent GWAS results excluding the HRS samples (Lee et al., 2018), from which SNP effects on years of education are obtained. Higher scores predict higher years of education and serve as indicators for a genetic predictor to educational attainment. The education PGS was standardised for the full sample so that effects can be interpreted as a ± 1 SD change relative to the sample. The relationship between education PGS, years of education and cognitive functions are presented in Supplementary Material Table S1.

The research method using genetic data may suffer from potential selection bias, as 262 263 respondents had to live until the 2006-2008 genotyping period. Of the original 37,495 264 respondents, 28,136 (75%) lived until at least 2006. Death of HRS participants prior to genotype collection in 2006, 2008 and 2010 may cause mortality selection bias. If individuals 265 with lower level of education PGS and worse cognition were more likely to die, the 266 association we estimated on the sample could be confounded (Domingue et al., 2017). To 267 268 alleviate the concern, we applied the inverse probability weighting to account for mortality 269 selection in our main analyses.

270 Covariates

Educational attainment is measured in years of education. We control for gender and population stratification for all analyses, as the frequencies of certain genetic variants vary by ancestral background. Ignoring genetic variation due to ancestry may result in population stratification bias when genetic effects are confounded by ancestry. Standard practice in accounting for population stratification using GWAS data is to include as covariates the first

few principal components that capture most of the genetic variation due to ancestry. We
adjusted for population stratification using the first ten principal components (Price,
Patterson, & Plenge, 2006).

279 Analytical strategy

Growth curve models were used to examine the individual cognitive trajectories of the respondents, which enabled us to study the effect of genetic predictor to educational attainment on the level of cognitive ability and its rate of change. We fit a linear, age-related decline random effect model and allow the age intercept and slope in the models to co-vary. Separate growth curve models were estimated with each cognitive measure as a dependent variable. Random effects included intercept and linear age, with the conventional unstructured covariance. A general specification of the model is

287
$$Cognition_{ij} = \beta_0 + \beta_1 \times (Age_j - \overline{Age}) + \beta_2 \times PGS_i + \beta_3 \times PGS_i \times (Age_j - \overline{Age}) + \beta_4 *$$

288
$$X_{ij} + \beta_5 * X_{ij} * (Age_j - \overline{Age}) + \mu_{ij} + \mu_{ij} * (Age_j - \overline{Age}) + \varepsilon_i$$

where age is centred around the grand mean (75), *Cognition*_{*ij*} represents the cognitive score for person i at age j, β_0 is the population mean of cognitive ability at the grand-mean age, β_1 represents the linear fixed effect of age, β_2 represents the effect of education PGS on the cognitive ability, β_3 is the linear effect of education PGS on the change rate of cognitive skills, β_4 and β_5 are the effects of X – a vector including individual covariates – on initial cognitive ability and the growth rate of change. σ_1 and σ_2 are the random intercept and slope. μ_{ij} and μ_{ij} are intercept and age variance.

296 **Results**

The rescaled cognitive scores represent the comparable percentage of correctly completed tasks. MS tasks were relatively easier compared to EM and A&C tasks. Older adults on average completed 80% and 70% of the A&C and MS tests respectively, while EM has only a mean score around 44, dragging GC towards 60 (<u>Table 1</u>). The mean trajectory in cognitive change over age is presented in <u>Figure 2</u>.

302 [Insert Table 1]

303 [Insert Figure 2]

304 Higher genetic predictor for education is associated with better cognitive performance,
305 independent of education.

306 <u>Figure 3</u> depicts the genetic effect sizes at age 75 from the growth curve models on each 307 cognitive measure. For each outcome, we explore two models: a model with education PGS 308 as the only predictor, and one with education PGS with education adjusted. Age, gender and 309 the first ten principal components are included in all the models.

310 [Insert Figure 3]

There is a clear pattern showing that education PGS are independently positively correlated with cognitive levels (Hypothesis 1). HRS respondents with a higher education PGS higher than their peers in cognitive tasks across all measures at age 75. The effect size of one standard deviation increase in education PGS on cognitive ability ranges from 1.9 to 5.7. Estimates are statistically significant (p < 0.001). Since education PGS and educational attainment are correlated ($\beta = 0.31$, p < 0.001), unsurprisingly the effect sizes drop after

education is controlled for, yet the effect sizes remain highly significant. After taking years of
education into account, the effect size of education PGS on EM, A&C, MS and GC declines
by 60%, 40%, 40% and 50%, respectively. These results indicate that education PGS
influence cognitive performance both independently, and through an education-mediated
pathway.

322 The effect of genetics on cognitive decline varies over age and by domains

The genetic influence on rate of decline is modelled by intercepts and slopes of the growth 323 324 curve as functions of education PGS and covariates. Figure 4 displays the predicted agespecific cognitive scores based on the fixed effects of education PGS (with and without 325 326 controlling for education). Education PGS is negatively associated with EM, and therefore a faster rate of decline ($\beta = -0.04$, p < 0.01). The effects indicate that higher education PGS 327 would lead to a faster rate of EM decline in old age. Individuals with higher education PGS 328 scores higher on GC and EM at the late stage of middle age, but the genetic effect diminishes 329 330 with age. This result contradicts our hypothesis 3, in that the advantage of a higher education PGS on GC and EM fades at old ages. For crystallised intelligence, in line with hypothesis 2, 331 332 higher education PGS does not change the rate of cognitive decline. Again, after controlling for education, the association between education PGS and the rate of cognitive decline 333 334 weakens.

335 [Insert Figure 4]

For GC, we found that in the education-unadjusted model, education PGS does not have a significant effect on GC decline. Surprisingly, when both education PGS and educational attainment are included in the model, the effect of education PGS becomes stronger and significant at the 0.01 level. education PGS is associated with a faster GC decline driven by 16

EM. The GC results indicate a suppression effect between education PGS and educational
attainment that statistical removal of the education PGS effect could increase the magnitude
of the relationship between years of education and cognitive decline.

343 We further examined whether the effect of education PGS could be mediated or confounded 344 by other covariates. We add social engagement, drinking, smoking and health conditions 345 individually to the education-adjusted models. A final full model includes all the covariates. Intercept results for education PGS and years of education are presented in Figure 5a. For 346 intercept, the effects of education PGS on cognitive performance does not change after 347 adjusting for covariates across all measured cognitive sub-domains. For slope, only education 348 PGS robustly predicts a faster rate of EM decline (Figure 5b). For general cognition, we 349 found that the effect of EA3 becomes insignificant on the rate of decline after including 350 smoking and pre-existing health conditions. 351

352 [Insert Figure 5]

353 Sensitivity analyses

We conducted sensitivity analyses to evaluate the consistency of findings. Details are 354 355 presented in the supplementary materials. First, to examine whether our results are driven by mortality selection, we compared our main analyses with models unadjusted for inverse 356 357 probability weights. The results from unweighted and weighted models are very similar. Weighted models improve the model fit measured by AIC and BIC. Further, we estimate our 358 models in four birth cohorts (before 1917, 1917-1926, 1927-1936, after 1937). However, the 359 360 association between education PGS and rate of change in EM loses its significance in every cohort, but the sign remains negative. This may indicate a lack of power from the small 361 sample for each cohort, as the sample size ranges from 859 to 2,485. Therefore, even though 362

weighted results reassure us that selection did not produce much bias, we cannot completelyrule out the competing explanation.

Second, the nature of the survey-based assessments may produce measurement error in 365 cognitive domains. We plot the coefficient of variation (standard deviation/mean) as an 366 367 indicator of measurement error (see Supplementary Material). It shows that the coefficient of variation increases with age slightly and becomes fairly unstable after age 90. We excluded 368 369 the respondents age 90 and above and ran our models again, and our conclusion holds after removing the oldest old. In addition, since cognitive measures are the dependent variables, 370 any measurement error is not likely to bias the estimated effect of education PGS but to 371 372 reduce the power of the statistical model. Our findings of lack of association hence should be 373 interpreted with caution.

Third, recent studies find that people with higher education PGS are more likely to be born in
socially advantaged families (Belsky et al., 2016; Belsky et al., 2018; Domingue et al., 2015).
Our results are robust after controlling for parental education as a measure of family origin.

Fourth, we control for the general cognition related polygenic score based on Davies et al. (2015). We examine whether education PGS influence cognitive performance via cognitionrelated genetic mechanisms. The magnitudes of estimates are slightly reduced, suggesting that education PGS predict cognitive performance and decline independently of cognitionlinked genetics. The effect of education PGS on each cognitive domain holds even after controlling for covariates, suggesting that genetic effects are not completely mediated by educational attainment and other mediators.

Finally, Keller (2014) has expressed scepticism on the positive findings from geneenvironment interaction studies in that potential confounders are not properly accounted for

in the statistical models used to test G×E effects. Including the potential confounders as covariates alone in the models may not be sufficient, as this practice does not control for the effects these variables might have on the gene-environment interaction. To show that the results in this study are robust after properly controlling for confounders, we re-ran the G×E models adding the covariate-by-environment (C×E) and the covariate-by-gene (C×G) interaction terms. The results are similar to the main analyses (see supplementary materials).

392 Discussion

In this study, we aim to explain the interpersonal variability in age-related cognitive decline 393 with education PGS. Existing research predominantly focuses on quantifying genetic and 394 395 environmental components of variance in cross-sectional cognitive data and has provided 396 evidence of genetic influences on cognitive ability (Davies et al., 2016; Rietveld & Webbink, 2016). Yet, few researchers have examined longitudinal cognitive change and genetic 397 predictor. Genes are inherited pre-birth and remain the same over a lifespan, but genetic 398 399 effects on phenotypes can vary over age as a function of gene expression associated with developmental timing or environmental circumstances (Lee, Gatz, Pedersen, & Prescott, 400 401 2016). Research to date has not offered information on changes in the genetic contribution to 402 individual heterogeneity in cognitive performance in older age.

403 Our main research question is whether education PGS is associated with higher initial level 404 and variation in cognitive abilities at the early stages of older adulthood. We analysed data on 405 the trajectory of cognitive performance across three individual and one aggregate domains in 406 over 5,000 individuals interviewed longitudinally as part of the HRS. In line with previous 407 literature, we find that education PGS predict a higher initial level of cognitive performances 408 independent of observed years of education, parental education, cognition-related PGS, and 409 other social factors. Our results on the cognitive decline are unlikely to be driven by selection 19

bias. In terms of the rate of cognitive change, the effect of education PGS on episodic
memory diminishes over age. We observe no association between education PGS and the rate
of change in the attention & concentration and mental status.

413 Results across a range of cognitive domains suggest that the education PGS is related to 414 significantly higher cognitive functions. Even after controlling for observed years of education, the relationship between education-associated genetic variants and cognitive 415 416 ability persists. The magnitude of the genetic effect size decreases in education adjusted models. These results are consistent with the evidence from Okbay et al. (2016), Rietveld et 417 al. (2013), and Rietveld et al. (2014), which suggests that there is an education PGS influence 418 on cognitive ability via both biological pleiotropy and mediated pleiotropy. The genetic 419 420 variants are associated with a particular neurotransmitter pathway involved in synaptic plasticity, which is the main cellular mechanism for learning and memory (Rietveld et al., 421 422 2014).

423 The analyses of cognitive trajectories caused by normal ageing showed that education PGS is related to the rate of cognitive decline, but the effect is only on episodic memory – a type of 424 fluid intelligence – and driving the same effect on global cognition. Performances in global 425 cognition and episodic memory are better in groups with higher education PGS for those 426 427 under 85 years old; this difference is completely attenuated over the age of 90 due to faster 428 cognitive decline in the high education PGS group. The findings on cognitive decline are in 429 agreement with recent studies showing that genetic effects vary in cognition with age (Lee et 430 al., 2016). However, the results contradict recent candidate gene analyses, which supports the 431 resource-modulation hypothesis (Laukka et al., 2013; Li et al., 2013; Papenberg et al., 2014). Candidate genes research focusing on a small amount of genetic variants may find a 432 magnifying effect during the ageing process via very specific biological channels (for 433 20

434 example, APOE influences memory through low-density lipopropotion cholesterol, high-435 density lipopropotion cholesterol, and tryglycerides) (Taylor et al., 2011). Such an effect is age-specific. Taylor et al. (2011) report a lack of association between APOE and cognitive 436 437 function in children. Belsky et al. (2016) adopt a polygenic score approach using growth curve modelling and finds that children with higher polygenic scores performed better on 438 439 cognitive tests and exhibited a faster pace of cognitive development during childhood. Their result, along with our findings, may suggest that education PGS are more important during 440 younger age, helping individuals to achieve higher education levels, but the protective effect 441 diminishes on episodic memory during the ageing process. Note that our analyses only 442 443 examine episodic memory as fluid intelligence due to data availability. Future research needs 444 to test more cognitive functions in order to generalise results to other types of fluid intelligence. Future research should also test cognitive change across a longer life span that 445 covers childhood, younger and middle adulthood to to comprehensively infer the 446 heterogeneity of genetic influence on the cognitive trajectory. 447

For global cognition, when we model education PGS and educational attainment separately, 448 449 both education PGS and education do not have any effect on the rate of cognitive decline. 450 When education PGS and years of education are jointly included in the model, education PGS and years of education both become statistically significant with opposite but more 451 452 substantial magnitudes of effects. This finding indicates that the education PGS and 453 phenotype confound each other via a suppression effect. Failure to take genetic predictor into 454 account may underestimate the protective effect from years of education, and the adverse 455 effect of genes for education.

456 Our study suffers from three main limitations. First, the variability in genetic effect may be457 due to ceiling and floor effects inherent in cognitive measures that narrow the potential range

458 of decline. Mental status as a crystallised intelligence tends to start declining at a later age 459 compared to fluid intelligence and is most pronounced in older adults with pathological brain damages (Albert, 1995). The finding that older adults with lower level of genetic predictor to 460 461 educational attainment experience a more rapid cognitive decline (compared to a more gradual decline for those with higher education PGS) could be due to ceiling effects in the 462 463 measurement that limit the variability of change for well-educated older adults with higher initial scores. People with higher education PGS thus enjoy higher cognition for their entire 464 adult life. More sensitive measures that cover greater variability in cognitive function might 465 provide more accurate estimates in future research. Sensitivity analyses excluding the 466 467 individuals who score the lowest 5% in each measure retained similar results, suggesting 468 floor effects do not compromise the analysis.

469 Second, although the polygenic score approach is superior to the traditional candidate genes 470 approach in many ways as mentioned above, it is not without limitation. The polygenic socre is based on mostly homogeneous groups of non-Hispanic Caucasian older adults in the U.S. 471 Our findings may not extend to individuals of other ethnic or cultural backgrounds, or later-472 473 born cohorts. Furthermore, the education PGS we use explains only a small proportion of Lee 474 et al.'s (2018) estimated genetic influence on educational attainment (Supplementary 475 Material). The genetic discoveries on education PGS do not account for gene-gene 476 interactions or gene-environment interactions. This may lead to measurement error in the 477 score. Our estimates may be thus biased toward zero (Conley et al., 2016), which provides a potential explanation for the lack of association between education PGS and the rate of 478 479 change in crystallised intelligence.

480 Despite its limitations, this study provides an essential contribution to existing knowledge on
 481 the variability of cognitive decline by genetics. Our results are consistent with recent research
 22

482 showing that education and cognitive ability are genetically correlated (Belsky et al., 2016; 483 Belsky et al., 2018; Wedow et al., 2018). We provide evidence that the causal link between educational attainment and cognitive abilities is subject to genetic confounding. Genetic 484 485 effects on cognition are not fully mediated by education and independent genetic influences 486 may exist in the relationship between education and cognitive decline. The associations 487 between a genetic predictor to educational attainment and cognitive decline that have been identified are especially relevant because they help to clarify the contributions of observed 488 489 education and genes to cognitive ageing. Future research should also consider genetic effects when investigating non-genetic factors in cognitive decline. Controlling for genetic effects 490 491 can avoid omitted variable bias when estimating environmental factors. The finding that the 492 genetic effect on cognitive decline for episodic memory decreases with age represents a need to understand the mechanisms between genetic endowment of educational attainment and 493 494 cognitive decline biological from angle. a

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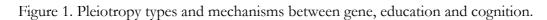
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Variables	Mean (SD) or Percentage
Outcomes: Cognitive Functions (rescaled)	Y
Episodic Memory (EM)	48.40 (16.67)
Attention and Concentration (A&C)	75.92 (29.61)
Mental Status (MS)	88.24 (13.34)
General Cognition (GC)	65.47 (12.69)
Exposure:	
Education PGS (Unstandardized)	-0.23 (0.14)
.ge	74.59 (6.99)
Gender (female)	57.56%
Years of Education (Unstandardized)	12.96 (2.52)
Social engagement	
Low	85.52%
Moderate	12.47%
High	2.01%
Current Smoker	8.31%
Drinking	
Non-Drinker	64.28%
Moderate-Drinker	34.85%
Heavy-Drinker	0.84%
Chronic Conditions	
No Condition	33.48%
1-2 Conditions	57.74%
More than 3 Conditions	8.78%



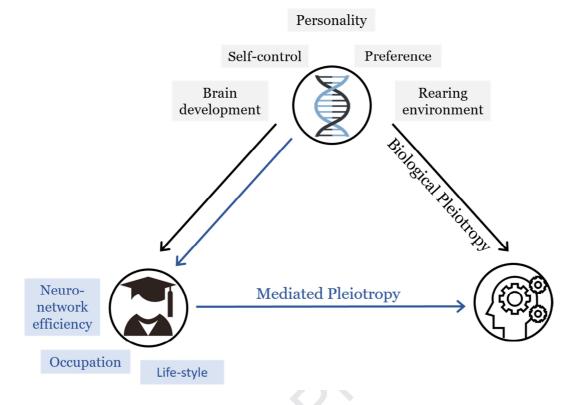


Figure 2. Box plots of the cognitive abilities over age groups (with outliers).

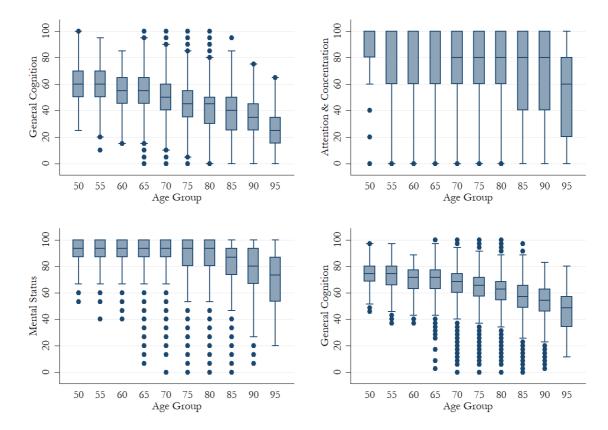
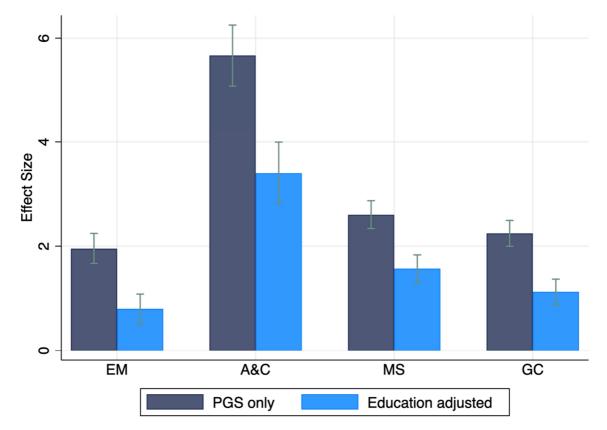
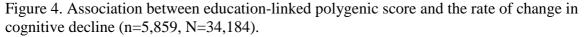
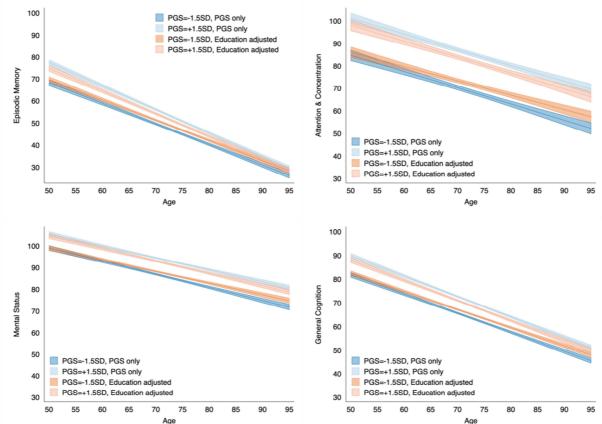


Figure 3. Association between education-linked polygenic score and level of cognitive abilities (n=5,859, N=34,184)



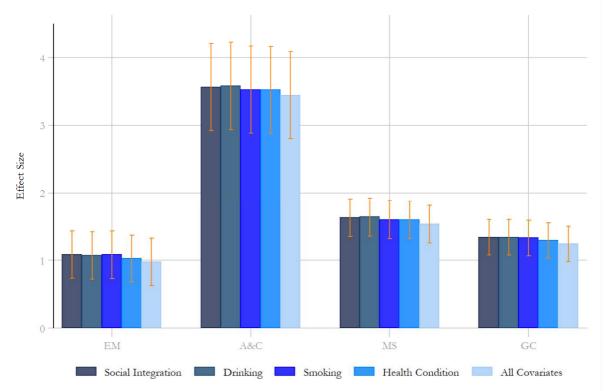
Note: The barplot shows the magnitude of the effect of a 1 SD increase in polygenic score on percentage correctness cognitive performance at age 75. Error bars indicate 95% confidence intervals. Dark navy bars show the effect sizes for the base models that education-linked polygenic score is the only predictor. Blue bars indicate the effect sizes for the education adjusted models. Both polygenic score and years of education are standardised. Gender and ten principal components are adjusted in all models. For details, see supplementary materials.



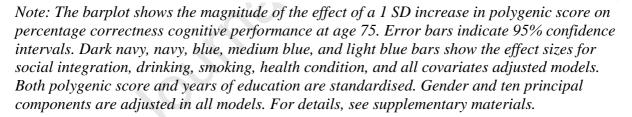


Note: The margin plots show the effect of a 1 SD increase in polygenic score on the rate of change in cognitive decline with age. The y-axes represent the percentage of correctness in the completed tasks for the domains. The shaded areas show 95% confidence intervals. Blue lines show the trajectories for the base models that education-linked polygenic score is the only predictor. Orange lines indicate the trajectories for the education adjusted models. Both polygenic score and years of education are standardised. Gender and ten principal components are adjusted in all models. For details, see supplementary materials.

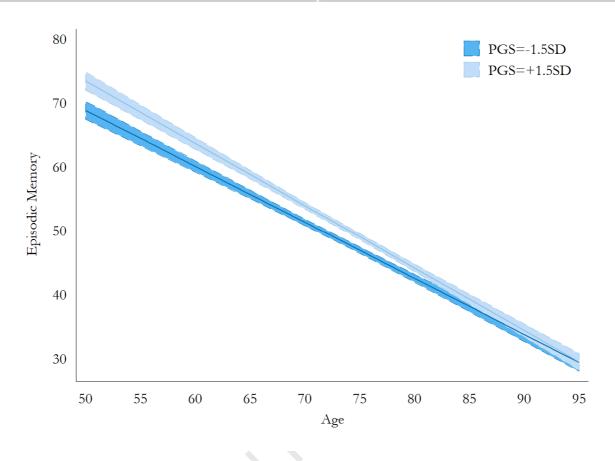
Figure 5. Intercept and slope results from the growth curve models on cognitive outcomes, controlling for covariates (n=5,859, N=34,184).



a). Intercept results for four cognitive outcomes.



b). Slope result for Episodic Memory, all covariates adjusted.



Note: The margin plots show the effect of a 1 SD increase in polygenic score on the rate of change in EM with age. The y-axes represent the percentage of correctness in the completed tasks for EM. The areas between the dashed lines show 95% confidence intervals. Both polygenic score and years of education are standardised. Gender and ten principal components are adjusted. For details, see supplementary materials.

- Older adults with higher scores perform better across all measures of cognition.
- The relationship is robust after controlling for phenotypic educational attainment.
- The genetic effect on episodic memory diminishes with age.
- Future research should consider genetic effects when examining cognitive decline.

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