

ORIGINAL ARTICLE

Intracluster correlation coefficients from school-based cluster randomized trials of interventions for improving health outcomes in pupils

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Abstract

Background and Objectives: To summarize intracluster correlation coefficient (ICC) estimates for pupil health outcomes from school-based cluster randomized trials (CRTs) across world regions and describe their relationship with study design characteristics and context.

Methods: School-based CRTs reporting ICCs for pupil health outcomes were identified through a literature search of MEDLINE (via Ovid). ICC estimates were summarized both overall and for different categories of study characteristics.

Results: Two hundred and forty-six articles reporting ICC estimates were identified. The median (interquartile range) ICC was 0.031 (0.011 to 0.08) at the school level ($N = 210$) and 0.063 (0.024 to 0.1) at the class level ($N = 46$). The distribution of ICCs at the school level was well described by the beta and exponential distributions. Besides larger ICCs in definitive trials than feasibility studies, there were no clear associations between study characteristics and ICC estimates.

Conclusion: The distribution of school-level ICCs worldwide was similar to previous summaries from studies in the United States. The description of the distribution of ICCs will help to inform sample size calculations and assess their sensitivity when designing future school-based CRTs of health interventions. © 2023 The Author(s). Published by Elsevier Inc. This is an open access article under the CC BY license (<http://creativecommons.org/licenses/by/4.0/>).

Keywords: Children; Cluster randomized trials; Intracluster correlation coefficient; Public health; Randomized trials; Schools

1. Background

Cluster randomized trials (CRTs) are studies in which clusters (groups) of individuals are randomized to trial arms and outcomes are measured on individuals [1]. CRTs are increasingly undertaken in schools to evaluate public health

interventions for improving outcomes of children and adolescents [2–5]. Schools provide a natural environment in which to recruit and study children and deliver interventions to improve their health due to the amount of time they spend there [3,6,7]. CRTs may be undertaken in schools because many of the interventions examined in such studies

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What is new?**Key findings**

- Few studies outside the United States (US) have summarized intracluster correlation coefficients (ICC) for pupil health outcomes and explored their size in relation to design characteristics in school-based cluster randomized trials (CRTs).

What this adds to what was known?

- This study collated 260 ICCs for school-related clusters from CRTs worldwide to inform sample size calculation for future trials.

What is the implication and what should change now?

- Two-thirds of school-level ICCs were no greater than 0.05 and three-quarters were under 0.08.
- The ICC distribution was similar to previous summaries from US-based studies and larger for definitive trials than feasibility studies.
- There was little evidence of relationships between ICC estimates and region, health outcome area, and educational level.

are designed to be delivered to entire schools or classrooms [4], interventions are theorized to affect change at those levels, and randomizing clusters (for example, schools, classes) helps to minimize contamination between trial arms that may otherwise occur if individuals are allocated [1,6,7].

CRTs require more participants than individually randomized trials because observations on individuals in the same cluster are usually more similar than those from different clusters [1]. Due to this lack of independence between individuals within clusters, if standard sample size formulae are used this may result in an underpowered study [1]. Correlation between pupils within clusters needs to be accounted for when designing and analyzing data from CRTs. In the sample size calculation, this is done by inflating the number of participants required in an individually randomized trial by the design effect (DE):

$$DE = 1 + (\bar{n} - 1)\rho$$

where \bar{n} is the mean number of participants providing outcome data in each cluster (cluster size) and ρ is the intracluster correlation coefficient (ICC) of the outcome [1]. The ICC quantifies the similarity of observations on individuals within clusters. For continuous outcomes, it can be defined as the proportion of the total variability in the outcome that is between clusters as opposed to between individuals within clusters:

$$ICC = \frac{\sigma_b^2}{\sigma_b^2 + \sigma_w^2}$$

where σ_b^2 is the between-cluster variance component and σ_w^2 is the within-cluster variance component [1]. Under this definition the ICC can take values between zero and one. The larger the ICC, the greater the sample size required. Similarity between participants from the same cluster can also be quantified by the between-cluster coefficient of variation (CV) of the outcome (the ratio of the between-cluster standard deviation to the outcome mean [6]):

$$CV = \frac{\sigma_b}{\mu}$$

where σ_b is the between-cluster standard deviation and μ is the mean outcome across the clusters [6]. The CV can then be incorporated into a modified design effect formula.

In the context of school-based CRTs there are several reasons for the similarity of outcomes between pupils within schools. First, in some countries, pupils and their parents/guardians have some influence regarding the school they attend [8]. Schools are likely to attract pupils with similar characteristics and who are more likely to share similar behaviours [3]. Second, pupils interact in the school setting and may influence the behaviour of their peers in the same schools or classrooms [8]. Finally, the school itself can influence the behaviours of pupils through its physical environment, ethos and policies [9,10].

At the time of sample size calculation the ICC is usually unknown and specification of a suitable value for the outcome and type of cluster should be informed by the empirical literature [1]. Researchers have reported ICCs for pupil health outcomes to be generally smaller than those for educational outcomes in schools [11–13]. This might be expected given that the main purpose of schools is to provide education [8]. Although ICCs for health outcomes in health care settings are well established, particularly in primary care [1,14,15], there is a relative lack of reported estimates in the school setting. Several studies have provided estimates of ICCs from school-based CRTs or surveys for outcomes related to substance use [8,16–24], nutrition [25–27], physical activity [24,27–29], and mental health and behaviour [12,24,30], but the vast majority of these were undertaken in the United States. It is not known whether these estimates are transferable to other regions and education systems, and outcome areas such as infectious diseases and dental health are not well represented. Furthermore, although patterns in the size of the ICC have been investigated [14,15,31–34], little is known about the extent to which ICCs from school-based CRTs differ by study characteristics.

A summary of ICCs for a range of health outcomes in different settings would aid the design of future school-based CRTs by providing plausible values that can be used

in sample size calculations. Estimates from CRTs specifically, rather than surveys, are potentially especially relevant as they may better reflect the level of variation in outcomes across the types of schools that tend to participate in health-related trials [1] (p177).

1.1. Objectives

This paper collates and summarizes ICC estimates for health outcomes from school-based CRTs and examines the relationship between the size of the ICC and study characteristics.

2. Methods

2.1. Data sources and search methods

A systematic searching approach was used to identify papers reporting ICC estimates from school-based CRTs. MEDLINE (Ovid) was exclusively searched for published peer-reviewed articles reporting school-based CRTs from inception to 18th October 2021. The search strategy was developed based on a strategy by Taljaard and colleagues [35] used to identify CRTs, combined with school-related terms (Table 1).

2.2. Inclusion and exclusion criteria

Eligible articles reported school-based studies with a CRT design, including articles reporting baseline data, follow-up outcomes, or secondary data analyses that used the data to address additional questions that were unrelated to the main trial objectives. To be eligible, the article had to report the estimate of an ICC/CV for at least one health outcome measured on pupils. The eligible study population was pupils attending pre-primary, primary, lower secondary

and higher secondary educational settings according to the United Nations Educational, Scientific and Cultural Organisation (UNESCO) International Standard Classification of Education (ISCED) system [36]. Eligible clusters were any school-related unit (e.g., schools, classes/classrooms, year groups, teachers). Any intervention(s) were considered. Articles were excluded if they randomised after-school clubs, school-based health centres or childcare centres. Articles that only reported protocol/design information, process evaluations, economic evaluations/cost-effectiveness analyses, statistical analysis plans, commentaries and mediation/mechanism analyses were also excluded.

If more than one publication of the same eligible study was identified, the key study report (index paper) for data extraction was determined by identifying the article that first published the outcomes.

2.3. Sifting and validation

Titles and abstracts were screened by two independent researchers (KP & OU) for eligibility against the inclusion criteria. Any studies for which the reviewers were uncertain of inclusion status were progressed to full text screening. Two independent researchers (KP & OU) examined the full text of each article against the inclusion criteria. Any disagreements over inclusion were resolved through discussion with a third researcher (MN).

2.4. Data extraction

One researcher (KP) extracted data from all included articles, while a second (OU) independently validated the process. Any uncertainty regarding the data extraction was resolved through discussion, or consultation with a third researcher (MN). The information extracted is specified in Table 2.

The ICC/CV estimate(s) of one pupil health outcome was extracted from each article, as estimates for multiple outcomes from the same study would likely be correlated and contribute relatively little additional information to the analyses in this paper which are focussed on comparing the ICC/CV across different study scenarios. Where estimates were reported for the chosen outcome at multiple levels (for example, school and class) these were all extracted. The criteria used to select the ICC/CV when multiple estimates were reported for a given paper are presented in Table 3. Where studies reported both unadjusted and adjusted ICCs, the former was extracted on the basis that this would be of more general use to future researchers who may want to adjust their estimate of the intervention effect for a specific set of prognostic factors. Where the ICC for a given outcome was reported for multiple time points the ICC for the earliest wave was extracted, as the ICC estimate would be less likely to be impacted by the intervention. For a similar reason, where

Table 1. Search strategy using MEDLINE (through Ovid)

Search strategy
Terms for Randomized Controlled trials:
1. random:.mp.
2. trial.ab, kw, ti.
Cluster design-related terms:
3. "cluster*".ab, kw, ti.
4. "communit*".ab, kw, ti.
5. group*adj2 random*.ab, kw, ti.
6. 3 OR 4 OR 5
School terms:
7. exp Schools/
8. School*.ab, kw, ti.
9. 7 OR 8
Final search stages:
10. 1 AND 2 AND 6 AND 9
11. 10 limited to English language

Table 2. Data extracted

Aspect	Information extracted
Publication details	Author surname, year of publication, title of article, type of study (that is, definitive or feasibility study).
Setting information	Country in which the study took place (for example, France), stage of education (for example, primary, secondary), gender of pupils, age(s) of pupils at baseline.
Study design	Type of cluster unit allocated, cluster unit of ICC/CV estimate.
Sample size information	ICC/CV assumed in the sample size calculation, number of clusters and pupils that provided outcome data, number of classes per school.
Health outcome information	Health area of outcome (for example, physical activity), outcome description (for example, amount of moderate-to-vigorous physical activity), outcome type (for example, continuous, binary), timing (months postrandomization) at which outcome was measured.
ICC information	ICC/CV of the outcome (and 95% CIs where provided), analytical method used to calculate ICC/CV (for example, multilevel model [37], marginal model using Generalized Estimating Equations [38]), whether the ICC/CV estimate was pooled across trial arms, whether the ICC/CV estimate was unadjusted or adjusted for prognostic factors, whether the ICC/CV estimate was adjusted for the baseline value of the outcome, whether the ICC/CV was estimated from an analysis of change scores between baseline and follow-up, whether a repeated measures analysis was used to estimate the ICC.

the ICC was reported separately for the control and intervention arms the former was chosen.

2.5. Data analysis

Study characteristics were summarized using medians, interquartile ranges (IQRs), and ranges for continuous variables and numbers and percentages for categorical variables. Mann–Whitney and Kruskal–Wallis tests were used to compare the ICC estimates across subgroups. Analyses were undertaken using Stata 17 [39].

3. Results

3.1. Search results

Three thousand six hundred and thirty-two articles were identified through searching MEDLINE. One thousand five hundred and ninety articles were included in the full text screening stage and 246 articles were identified as eligible for inclusion in the review. One paper reported an estimate of the between-cluster coefficient of variation of the outcome, but this was negative and therefore the paper

was not included. The PRISMA flow diagram is presented in [Figure 1](#).

3.2. Publication characteristics

Worldwide, the rate of publication of articles reporting ICC estimates from school-based CRTs that evaluate interventions for improving pupil health outcomes has increased since the first publication in 1999; 44 articles were published between 1999 and 2010, compared to 25 in 2021 alone. Of the 246 included studies, 226 (91.9%) were definitive trials and 20 (8.1%) were feasibility studies. The settings of included studies spanned all regions of the world and different stages of education. The majority of studies ($n = 227$; 92.3%) included males and females. In most of the studies schools were the units of randomization ($n = 220$; 89.4%); classes were randomized in 23 (9.3%) studies; and school buildings [40], student groups [41] and year groups [42] were randomized in one study each. The studies spanned a range of different health outcome areas, the most common being socioemotional functioning and its influences ($n = 53$; 21.5%), physical activity

Table 3. Criteria used to select which ICC/CV to extract

Aspect	Criteria
Outcome measure	In the first instance, the ICC/CV for the primary health outcome was selected. If there was more than one primary health outcome, the ICC/CV for the first primary outcome presented in the Results section of the paper was selected. If no primary health outcome was declared, the ICC/CV for the health outcome on which the sample size calculation was based was selected. If no primary health outcome was declared and the sample size was not based on a health outcome, the ICC/CV for the first health outcome reported in the Results section of the paper was selected.
Time point at which outcome was measured	In the first instance, the ICC/CV from the baseline time point was selected. If this was not reported, the ICC/CV from the earliest time point of measurement was selected.
Unadjusted vs. adjusted ICC/CV	If the study presented both unadjusted ICCs/CVs estimates and estimates that are adjusted for prognostic factors, the unadjusted ICC/CV was extracted.
Control verses intervention arm	If the ICC/CV was reported separately for the intervention and control arms, the ICC/CV from the control arm was selected.

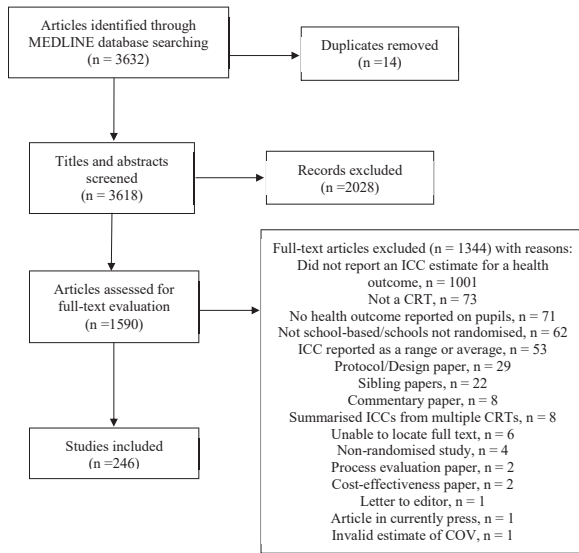


Fig. 1. PRISMA flowchart summarizing the results of the literature search and screening for eligibility.

($n = 34$; 13.8%), adiposity ($n = 28$; 11.4%) and smoking ($n = 21$; 8.5%).

Two hundred and sixty ICC estimates were extracted: 210 at school level, 46 at classroom level, and one each at the levels of school building [40], student group [41], year group [42] and sports-team [43]. Forty-five (17.3%) ICCs were estimated using the baseline measurement of the outcome. ICCs were extracted for 172 continuous outcomes, 78 binary outcomes, six count/rate outcomes, and two ordinal outcomes; for two extracted ICCs the outcome type was unclear. Of the studies that reported school-level ICCs, the median (IQR) number of clusters and pupils were 22 (12 to 40) and 1,110 (441 to 2,443), respectively. Of the studies that reported class-level ICCs, the median (IQR) number of clusters and pupils were 47 (25 to 88) and 647.5 (288 to 1,477), respectively.

3.3. Summary of ICC estimates

The median (IQR; range) school-level ICC estimate was 0.031 (0.011 to 0.08; 0 to 0.47); 51 (24.3%) of the school-level ICCs were less than or equal to 0.01 and 135 (64.3%) were less than or equal to 0.05. The mean (SD) school-level ICC was 0.060 (0.076). [Figure 2](#) summarizes the distribution of school-level ICCs. Both the beta distribution (with shape parameters 0.77 and 11.0) and the exponential distribution provided a good fit to the school-level ICC estimates. The median (IQR; range) of the class-level ICC estimates was 0.063 (0.024 to 0.1; -0.009 to 0.262); the only negative reported ICC was at this level. All ICC estimates are reported in [Appendix 1](#). School-level and class-level ICCs are reported side-by-side for 14 studies that reported at both those levels in [Appendix 2](#).

[Table 4](#) reports the median school-level ICC by region, health outcome area (for the 10 most common

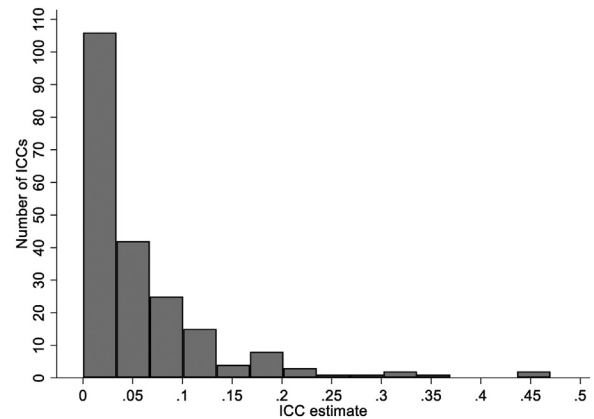


Fig. 2. The distribution of school-level ICCs in school-based CRTs ($N = 210$).

areas—there were at least eight ICC estimates for each area) and education stage. [Figure 3](#) uses dot plots to describe the distributions. Tests of significance indicated little evidence of differences across these subgroups and there was generally a fair amount of overlap in ICC distributions. The distribution of ICCs for the USA/Canada region (median 0.033 and 75% of estimates being lower than 0.073) is in keeping with summaries of USA-based estimates that have previously been reported [12,16–22,25–30]. There was reasonable overlap with the distributions in the other regions with the exception of Australia/New Zealand, for which the median and upper quartile were notably lower. The school-level ICC distributions for adiposity, physical activity and general health were lower than for other outcome areas. For two specific outcomes there were more than 10 estimates of the school-level ICC. For the 17 articles that reported the school-level ICC for body mass index (BMI) the median (IQR) was 0.021 (0.015 to 0.04) and for the 11 articles that reported amount of moderate-to-vigorous physical activity (MVPA) the median (IQR) school-level ICC was 0.018 (0.01 to 0.057). Although the median was higher, the overall distribution of ICCs was located at lower values for the preprimary stage compared to the later stages of education.

The median (IQR) school-level ICC was higher for definitive studies ($N = 192$) than feasibility studies ($N = 18$) (0.038 (0.016 to 0.08) vs. 0.01 (0.0005 to 0.04); $P = 0.005$). The median (IQR) school-level ICC was larger for continuous outcomes ($N = 135$) than binary outcomes ($N = 68$) although there was little evidence of a true difference in the distributions (0.04 (0.014 to 0.08) vs. 0.025 (0.008 to 0.08); $P = 0.21$). Summaries of the school-level ICCs are reported separately for continuous and binary outcomes in [Appendix 3](#).

For continuous outcomes, the median (IQR) school-level ICC was higher for studies that adjusted for the baseline of the outcome at the pupil level ($N = 35$) compared with those that did not ($N = 95$), but there was little evidence of a real difference (0.045 (0.013 to 0.09) vs. 0.040

Table 4. Median (IQR; range) school-level ICC by region, outcome area and education stage

Characteristic	N	Median ICC (IQR; range)	P value
Region			0.26
Europe ^a	45	0.04 (0.014 to 0.08; 0 to 0.47)	
USA and Canada	44	0.033 (0.010 to 0.073; 0 to 0.286)	
UK ^b	40	0.029 (0.01 to 0.106; 0 to 0.45)	
Australia and New Zealand	27	0.02 (0.01 to 0.03; 0 to 0.16)	
Asia ^c	21	0.05 (0.013 to 0.118; 0 to 0.31)	
Central and South America ^d	17	0.05 (0.016 to 0.09; 0.0001 to 0.36)	
Africa ^e	16	0.05 (0.018 to 0.127; 0.0005 to 0.21)	
Health outcome area			0.76
Socioemotional functioning and its influences ^f	39	0.05 (0.02 to 0.097; 0 to 0.217)	
Physical activity	30	0.035 (0.013 to 0.059; 0 to 0.19)	
Adiposity	26	0.027 (0.014 to 0.041; 0.004 to 0.19)	
Smoking	19	0.055 (0.017 to 0.11; 0 to 0.286)	
Alcohol use	10	0.055 (0.02 to 0.098; 0 to 0.121)	
Dental/oral health	10	0.051 (0.027 to 0.119; 0 to 0.31)	
General health	10	0.025 (0.014 to 0.045; 0.001 to 0.18)	
Infectious disease	9	0.042 (0.004 to 0.070; 0.0001 to 0.21)	
Nutrition	8	0.06 (0.010 to 0.097; 0 to 0.36)	
Violence	8	0.048 (0.014 to 0.085; 0.002 to 0.13)	
Education stage			0.40
Preprimary education only ^g	13	0.048 (0.03 to 0.063; 0 to 0.097)	
Primary education only ^h	81	0.04 (0.013 to 0.094; 0 to 0.47)	
Secondary education only ⁱ	81	0.03 (0.01 to 0.07; 0 to 0.31)	

^a Included countries stated as follows: Finland, The Netherlands, Denmark, Belgium, Norway, Germany, Estonia, Poland, Spain, Switzerland, Cyprus, Italy, Greece, Hungary, Sweden, Austria, Majorca, France, Ireland, Romania, Slovenia.

^b Included countries stated as follows: England, Northern Ireland, Scotland, Wales.

^c Included countries stated as follows: Israel, China, Iran, India, Japan, Bangladesh, Nepal, Taiwan, Peru, Pakistan, Thailand, Indonesia, Hong Kong.

^d Included countries stated as follows: Jamaica, Brazil, Ecuador, Chile, Haiti, Belize.

^e Included countries stated as follows: Uganda, South Africa, Kenya, Tanzania, Burundi.

^f Includes mental health, behaviour, neurodiversity, wellbeing, quality of life, bullying, social and emotional learning, body image, and self-esteem.

^g Includes preschools, kindergartens, educational childcare centres, and head-start schools.

^h Includes elementary schools, middle schools (Grade 6).

ⁱ Includes secondary schools, middle schools (\geq Grade 7), high schools, junior high schools, lower secondary schools, higher/upper secondary schools, vocational schools, intermediate vocational schools, secondary-level vocational schools, and continuation schools.

(0.016 to 0.07); $P = 0.50$). Also, for continuous outcomes, the median school-level ICC was identical for studies that did ($N = 11$) and did not ($N = 124$) analyze change scores (0.04; $P = 0.37$). The median (IQR) school-level ICC was lower for studies that estimated the ICC from a repeated measures analysis ($N = 37$) compared with those that did not ($N = 173$) (0.027 (0.01 to 0.057) vs. 0.036 (0.013 to 0.088)), but with little evidence of a systematic difference ($P = 0.15$). Finally, for binary outcomes, there was weak evidence that the median (IQR) ICC was higher for studies that use multilevel logistic regression to estimate this parameter on the logistic scale ($N = 42$) than those that use other methods to estimate it on the proportions (natural) scale ($N = 14$) (0.049 (0.014 to 0.109) vs. 0.014 (0.007 to 0.023); $P = 0.08$). The direction of this difference is consistent with the fact that the ICC on the logistic scale is generally larger than on the proportions scale [44]. [Appendix 4](#)

summarises the relationship between ICC estimates and the prevalence for binary outcomes.

4. Discussion

To our knowledge, this is the first paper to report the distribution of ICCs for pupil health outcomes from school-based CRTs worldwide. 260 ICC estimates from 246 school-based CRTs were extracted for outcomes spanning a range of health areas. There were few clear patterns regarding the relationship of the ICC with aspects of the design and analysis. Indeed, comparison of the ICC across categories of the study features examined was characterized by overlap in the distributions, although the differences in medians would be large in terms of the impact they would have on the sample size requirement for a CRT. Imprecision

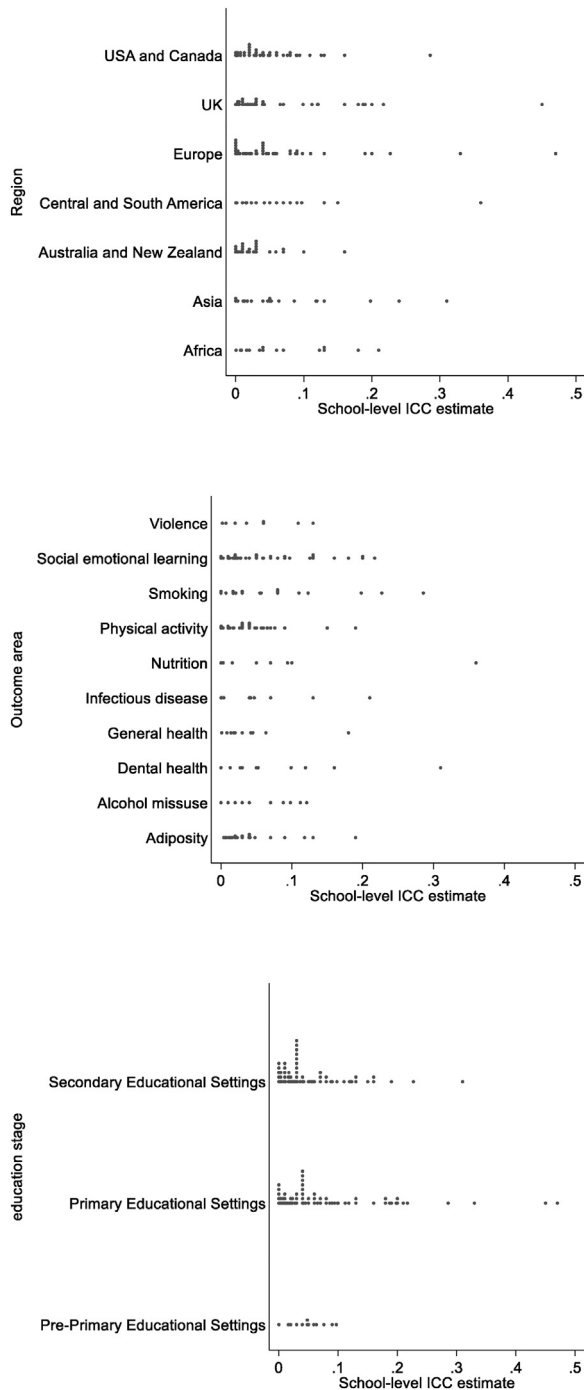


Fig. 3. Dot plots of school-level ICCs by region, outcome area, and education stage.

in the ICC estimates may have reduced the power to detect differences between subgroups defined by design and analysis characteristics.

The large number of different outcomes represented (Appendix 1) partly accounts for the variation in the estimates, although there was even a marked variation in ICC estimates across studies for the same outcome (that is, amount of MVPA and BMI). Sampling variability, the methodological context of the trials and the models

specified to estimate the parameter will also contribute to variability in the ICC estimates. Given the clinical and methodological heterogeneity across CRTs, an individual ICC estimate for a given outcome from a single study may have poor generalizability [28], and it has been recommended that researchers use the distribution of ICCs from many studies to model the sensitivity of sample size calculations [1,14,34]. Distributions of ICC estimates for health outcomes in primary care-based clusters have been found to be well described by the beta distribution [14,45]. The beta distribution was a good fit to the school-level ICCs reported in this paper as was the exponential distribution. The distribution parameters of these ICC estimates are of value for constructing informative priors when using a Bayesian framework to incorporate uncertainty about the ICC in sample size calculations for school-based CRTs [46,47].

There was little difference between ICC estimates that were adjusted for the baseline outcome measurement and those that were not. This may be due to differences in aspects of the design and setting across the studies and the fact that adjustment for individual-level prognostic factors may increase or decrease the ICC depending on the extent to which the between-cluster and within-cluster components of variance are reduced following adjustment [48].

The ICC from a repeated measures analysis using outcome data from across all study waves does not necessarily estimate the same parameter as an ICC for the outcome at a specific study wave. The correlation between observations from the same cluster from different waves may be smaller than the correlation between observations from the same cluster at the same study wave [34,49]. In this study, however, there was little evidence that the school-level ICC is lower for studies that estimate the ICC from a repeated measures analysis than those that do not, although the median was lower for the former set of studies.

Previously reported summaries of school-based ICCs for pupil health outcomes have largely used data from trials and surveys in the United States [12,16–22,25–30]. The distribution of school-level ICCs worldwide in the current paper was broadly similar to those previous summaries, with most estimates less than 0.05 and few greater than 0.1. Only the distribution for the Australia/New Zealand region was notably different (smaller).

The median ICC for pupil health outcomes was 0.031 at the school level and 0.063 at the class level. The difference is intuitive given the greater opportunity for interaction within classes as opposed to between classes within the same school and that the ICC has been reported to be larger when the natural cluster size is smaller [20,50]. The median ICC was markedly smaller for feasibility studies than in definitive trials. This may reflect that schools recruited in feasibility studies are a more restricted and less representative subset of the wider types of schools that are recruited in larger definitive studies [1] (p180/181). There was little evidence of a relationship between the ICC for pupil health

outcomes and stage of education. Previously, it has been reported that there is a tendency for ICCs for educational outcomes to be larger for lower education grades [48].

4.1. Strengths and limitations

This is the first study to collate and summarize ICCs for pupil outcomes across different health areas from school-based CRTs worldwide. The study used a systematic searching approach with dual screening and data validation. The sample of 246 CRTs was not sufficiently large to describe the ICC within different combinations of categories of the study design parameters (for example, only one combination of region and health outcome area provided at least 10 school-level ICC estimates). Partly for this reason, when investigating geographic variation in the ICC, we grouped countries into regions which will have obscured differences between individual countries. Based on empirical evidence from a European-based survey, it has been suggested that the ICCs assumed in the sample size calculation for school-based trials should be country-specific and outcome-specific [8]. As more school-based CRTs are undertaken the pool of reported ICCs will increase, enabling a more detailed examination with greater power to detect ICC patterns in relation to key study characteristics.

A potential limitation was the decision to use only the MEDLINE database. Although findings from a previous systematic review of similar studies indicated that few additional studies would have been found by searching other databases (specifically, EMBASE, DARE, PsychINFO, and ERIC) [4], we acknowledge that further articles may have been found by searching the grey literature. Additionally, some older articles may have been missed because the titles and abstracts did not refer to using a cluster design.

It was decided to extract the ICC estimate for only one outcome from each study even when multiple ones were reported. We anticipated that ICCs would be more similar within studies and wanted to avoid a scenario where a small number of studies that reported many ICCs had a disproportionate impact on the observed distribution of ICCs.

5. Conclusions

The 260 reported ICC estimates from studies spanning all world regions and different health outcome areas, and the summaries of their distribution are a valuable resource to researchers for calculating sample size for future school-based CRTs. The ICCs had a similar distribution to published summaries of the parameter from studies based in the United States. Better reporting of the ICC in CRTs, in keeping with CONSORT guidance [51], will provide a larger pool of data that can be used to explore the distribution of ICC values and the factors that determine them in greater detail.

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Supplementary data

Supplementary data to this article can be found online at <https://doi.org/10.1016/j.jclinepi.2023.03.020>.

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